

# Package ‘SANple’

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**Type** Package

**Title** Fitting Shared Atoms Nested Models via Markov Chains Monte Carlo

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**URL** <https://github.com/laura-dangelo/SANple>

**BugReports** <https://github.com/laura-dangelo/SANple/issues>

**Description** Estimate Bayesian nested mixture models via Markov Chain Monte Carlo methods. Specifically, the package implements the common atoms model (Denti et al., 2023), its finite version (D'Angelo et al., 2023), and a hybrid finite-infinite model.

All models use Gaussian mixtures with a normal-inverse-gamma prior distribution on the parameters. Additional functions are provided to help analyzing the results of the fitting procedure.

References:

Denti, Camerlenghi, Guindani, Mira (2023) <[doi:10.1080/01621459.2021.1933499](https://doi.org/10.1080/01621459.2021.1933499)>,

D'Angelo, Canale, Yu, Guindani (2023) <[doi:10.1111/biom.13626](https://doi.org/10.1111/biom.13626)>.

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estimate_clusters	<i>Estimate observational and distributional clusters</i>
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### Description

Given the MCMC output, estimate the observational and distributional partitions using `salso::salso()`.

### Usage

```
estimate_clusters(object, burnin = NULL, ncores = 0)
```

### Arguments

object	object of class SANmcmc (the result of a call to <code>sample_fiSAN</code> , <code>sample_fiSAN_sparsemix</code> , <code>sample_fSAN</code> , <code>sample_fSAN_sparsemix</code> , or <code>sample_CAM</code> ).
burnin	the length of the burn-in to be discarded before estimating the clusters (default is 2/3 of the iterations).
ncores	the number of CPU cores to use, i.e., the number of simultaneous runs at any given time. A value of zero indicates to use all cores on the system.

### Value

Object of class SANclusters. The object contains:

`est_oc` estimated partition at the observational level. It is an object of class `salso.estimate`.

`est_dc` estimated partition at the distributional level. It is an object of class `salso.estimate`.

`clus_means` cluster-specific sample means of the estimated partition.

`clus_vars` cluster-specific sample variances of the estimated partition.

### See Also

`salso::salso()`, `print.SANmcmc`, `plot.SANmcmc`, `print.SANclusters`

**Examples**

```

set.seed(123)
y <- c(rnorm(40,0,0.3), rnorm(20,5,0.3))
g <- c(rep(1,30), rep(2, 30))
out <- sample_fiSAN(nrep = 500, y = y, group = g,
                  nclus_start = 2,
                  maxK = 20, maxL = 20,
                  beta = 1)
estimate_clusters(out)

```

---

plot.SANmcmc

*Plotting MCMC output*


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**Description**

Plot method for objects of class SANmcmc. The function displays two graphs, meant to analyze the estimated distributional and observational clusters.

**Usage**

```

## S3 method for class 'SANmcmc'
plot(
  x,
  type = c("boxplot", "ecdf", "scatter"),
  estimated_clusters = NULL,
  burnin = NULL,
  palette_brewed = FALSE,
  ncores = 0,
  ...
)

```

**Arguments**

**x** object of class SANmcmc (the result of a call to [sample\\_fiSAN](#), [sample\\_fiSAN\\_sparsemix](#), [sample\\_fSAN](#), [sample\\_fSAN\\_sparsemix](#), or [sample\\_CAM](#)).

**type** what type of plot should be drawn (only for the left-side plot). Possible types are "boxplot", "ecdf", and "scatter".

**estimated\_clusters** the output of a call to [estimate\\_clusters](#) (optional). It can be used to speed up the function if the partition has already been computed. If `estimated_clusters = NULL`, the displayed partition is computed using [estimate\\_clusters](#).

**burnin** the length of the burn-in to be discarded (default is 2/3 of the iterations).

**palette\_brewed** (logical) the color palette to be used. Default is R base colors (`palette_brewed = FALSE`).

`ncores` if the partition is computed, the number of CPU cores to use to estimate the clusters, i.e., the number of simultaneous runs at any given time. A value of zero indicates to use all cores on the system.

`...` additional graphical parameters to be passed when `type = "scatter"` is used.

**Value**

The function plots a summary of the fitted model.

**See Also**

[print.SANmcmc](#), [estimate\\_clusters](#)

**Examples**

```
set.seed(123)
y <- c(rnorm(40,0,0.3), rnorm(20,5,0.3))
g <- c(rep(1,30), rep(2, 30))
out <- sample_fiSAN(nrep = 500, y = y, group = g,
                  nclus_start = 2,
                  maxK = 20, maxL = 20,
                  beta = 1)
plot(out, type = "ecdf", palette_brewed = TRUE)
```

---

`print.SANclusters`      *Print cluster summary*

---

**Description**

Print the cluster-specific sample means and variances of the estimated observational and distributional partition.

**Usage**

```
## S3 method for class 'SANclusters'
print(x, ...)
```

**Arguments**

`x` object of class `SANclusters` (the result of a call to [estimate\\_clusters](#))

`...` ignored.

**Value**

The function prints a summary of the estimated clusters.

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print.SANmcmc	<i>Print MCMC output</i>
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**Description**

Print method for objects of class SANmcmc.

**Usage**

```
## S3 method for class 'SANmcmc'
print(x, ...)
```

**Arguments**

x	object of class SANmcmc (the result of a call to <a href="#">sample_fiSAN</a> , <a href="#">sample_fiSAN_sparsemix</a> , <a href="#">sample_fSAN</a> , <a href="#">sample_fSAN_sparsemix</a> , or <a href="#">sample_CAM</a> ).
...	ignored.

**Value**

The function prints a summary of the fitted model.

**See Also**

[estimate\\_clusters](#), [plot.SANmcmc](#)

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sample_CAM	<i>Sample CAM</i>
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**Description**

`sample_CAM` is used to perform posterior inference under the common atoms model (CAM) of Denti et al. (2023) with Gaussian likelihood. The model uses Dirichlet process mixtures (DPM) at both the observational and distributional levels. The implemented algorithm is based on the nested slice sampler of Denti et al. (2023), based on the algorithm of Kalli, Griffin and Walker (2011).

**Usage**

```
sample_CAM(nrep, y, group,
           maxK = 50, maxL = 50,
           m0 = 0, tau0 = 0.1, lambda0 = 3, gamma0 = 2,
           hyp_alpha1 = 1, hyp_alpha2 = 1,
           hyp_beta1 = 1, hyp_beta2 = 1,
           alpha = NULL, beta = NULL,
           warmstart = TRUE, nclus_start = NULL,
```

```
mu_start = NULL, sigma2_start = NULL,
M_start = NULL, S_start = NULL,
alpha_start = NULL, beta_start = NULL,
progress = TRUE, seed = NULL)
```

## Arguments

nrep	Number of MCMC iterations.
y	Vector of observations.
group	Vector of the same length of y indicating the group membership (numeric).
maxK	Maximum number of distributional clusters (default = 50).
maxL	Maximum number of observational clusters (default = 50).
m0, tau0, lambda0, gamma0	Hyperparameters on $(\mu, \sigma^2) \sim NIG(m_0, \tau_0, \lambda_0, \gamma_0)$ . Default is (0, 0.1, 3, 2).
hyp_alpha1, hyp_alpha2	If a random $\alpha$ is used, (hyp_alpha1, hyp_alpha2) specify the hyperparameters. Default is (1,1). The prior is $\alpha \sim \text{Gamma}(\text{hyp\_alpha1}, \text{hyp\_alpha2})$ .
hyp_beta1, hyp_beta2	If a random $\beta$ is used, (hyp_beta1, hyp_beta2) specify the hyperparameters. Default is (1,1). The prior is $\beta \sim \text{Gamma}(\text{hyp\_beta1}, \text{hyp\_beta2})$ .
alpha	Distributional DP parameter if fixed (optional). The distribution is $\pi \sim GEM(\alpha)$ .
beta	Observational DP parameter if fixed (optional). The distribution is $\omega_k \sim GEM(\beta)$ .
warmstart, nclus_start	Initialization of the observational clustering. warmstart is logical parameter (default = TRUE) of whether a kmeans clustering should be used to initialize the chains. An initial guess of the number of observational clusters can be passed via the nclus_start parameter (optional). Default is nclus_start = min(c(maxL, 30)).
mu_start, sigma2_start, M_start, S_start, alpha_start, beta_start	Starting points of the MCMC chains (optional). mu_start, sigma2_start are vectors of length maxL. M_start is a vector of observational cluster allocation of length N. S_start is a vector of observational cluster allocation of length J. alpha_start, alpha_start are numeric.
progress	show a progress bar? (logical, default TRUE).
seed	set a fixed seed.

## Details

### Data structure

The common atoms mixture model is used to perform inference in nested settings, where the data are organized into  $J$  groups. The data should be continuous observations  $(Y_1, \dots, Y_J)$ , where each  $Y_j = (y_{1,j}, \dots, y_{n_j,j})$  contains the  $n_j$  observations from group  $j$ , for  $j = 1, \dots, J$ . The function takes as input the data as a numeric vector  $y$  in this concatenated form. Hence  $y$  should be a vector of length  $n_1 + \dots + n_J$ . The group parameter is a numeric vector of the same size as  $y$  indicating the group membership for each individual observation. Notice that with this specification

the observations in the same group need not be contiguous as long as the correspondence between the variables  $y$  and  $group$  is maintained.

### Model

The data are modeled using a univariate Gaussian likelihood, where both the mean and the variance are observational-cluster-specific, i.e.,

$$y_{i,j} | M_{i,j} = l \sim N(\mu_l, \sigma_l^2)$$

where  $M_{i,j} \in \{1, 2, \dots\}$  is the observational cluster indicator of observation  $i$  in group  $j$ . The prior on the model parameters is a Normal-Inverse-Gamma distribution  $(\mu_l, \sigma_l^2) \sim NIG(m_0, \tau_0, \lambda_0, \gamma_0)$ , i.e.,  $\mu_l | \sigma_l^2 \sim N(m_0, \sigma_l^2/\tau_0)$ ,  $1/\sigma_l^2 \sim Gamma(\lambda_0, \gamma_0)$  (shape, rate).

### Clustering

The model performs a clustering of both observations and groups. The clustering of groups (distributional clustering) is provided by the allocation variables  $S_j \in \{1, 2, \dots\}$ , with

$$Pr(S_j = k | \dots) = \pi_k \quad \text{for } k = 1, 2, \dots$$

The distribution of the probabilities is  $\{\pi_k\}_{k=1}^{\infty} \sim GEM(\alpha)$ , where GEM is the Griffiths-Engen-McCloskey distribution of parameter  $\alpha$ , which characterizes the stick-breaking construction of the DP (Sethuraman, 1994).

The clustering of observations (observational clustering) is provided by the allocation variables  $M_{i,j} \in \{1, 2, \dots\}$ , with

$$Pr(M_{i,j} = l | S_j = k, \dots) = \omega_{l,k} \quad \text{for } k = 1, 2, \dots; l = 1, 2, \dots$$

The distribution of the probabilities is  $\{\omega_{l,k}\}_{l=1}^{\infty} \sim GEM(\beta)$  for all  $k = 1, 2, \dots$

### Value

sample\_CAM returns four objects:

- **model**: name of the fitted model.
- **params**: list containing the data and the parameters used in the simulation. Details below.
- **sim**: list containing the simulated values (MCMC chains). Details below.
- **time**: total computation time.

**Data and parameters**: **params** is a list with the following components:

**nrep** Number of MCMC iterations.

**y, group** Data and group vectors.

**maxK, maxL** Maximum number of distributional and observational clusters.

**m0, tau0, lambda0, gamma0** Model hyperparameters.

**(hyp\_alpha1, hyp\_alpha2) or alpha** Either the hyperparameters on  $\alpha$  (if  $\alpha$  random), or the value for  $\alpha$  (if fixed).

**(hyp\_beta1, hyp\_beta2) or beta** Either the hyperparameters on  $\beta$  (if  $\beta$  random), or the value for  $\beta$  (if fixed).

**Simulated values:** `sim` is a list with the following components:

`mu` Matrix of size  $(nrep, maxL)$ . Each row is a posterior sample of the mean parameter for each observational cluster  $(\mu_1, \dots, \mu_L)$ .

`sigma2` Matrix of size  $(nrep, maxL)$ . Each row is a posterior sample of the variance parameter for each observational cluster  $(\sigma_1^2, \dots, \sigma_L^2)$ .

`obs_cluster` Matrix of size  $(nrep, n)$ , with  $n = \text{length}(y)$ . Each row is a posterior sample of the observational cluster allocation variables  $(M_{1,1}, \dots, M_{n,J})$ .

`distr_cluster` Matrix of size  $(nrep, J)$ , with  $J = \text{length}(\text{unique}(\text{group}))$ . Each row is a posterior sample of the distributional cluster allocation variables  $(S_1, \dots, S_J)$ .

`pi` Matrix of size  $(nrep, maxK)$ . Each row is a posterior sample of the distributional cluster probabilities  $(\pi_1, \dots, \pi_{maxK})$ .

`omega` 3-d array of size  $(maxL, maxK, nrep)$ . Each slice is a posterior sample of the observational cluster probabilities. In each slice, each column  $k$  is a vector (of length  $maxL$ ) observational cluster probabilities  $(\omega_{1,k}, \dots, \omega_{maxL,k})$  for distributional cluster  $k$ .

`alpha` Vector of length  $nrep$  of posterior samples of the parameter  $\alpha$ .

`beta` Vector of length  $nrep$  of posterior samples of the parameter  $\beta$ .

`maxK` Vector of length  $nrep$  of the number of distributional DP components used by the slice sampler.

`maxL` Vector of length  $nrep$  of the number of observational DP components used by the slice sampler.

## References

Denti, F., Camerlenghi, F., Guindani, M., and Mira, A. (2023). A Common Atoms Model for the Bayesian Nonparametric Analysis of Nested Data. *Journal of the American Statistical Association*, 118(541), 405-416. <doi:10.1080/01621459.2021.1933499>

Kalli, M., Griffin, J.E., and Walker, S.G. (2011). Slice Sampling Mixture Models, *Statistics and Computing*, 21, 93-105. <doi:10.1007/s11222-009-9150-y>

Sethuraman, A.J. (1994). A Constructive Definition of Dirichlet Priors, *Statistica Sinica*, 4, 639-650.

## Examples

```
set.seed(123)
y <- c(rnorm(40,0,0.3), rnorm(20,5,0.3))
g <- c(rep(1,30), rep(2, 30))
plot(density(y[g==1]), xlim = c(-5,10))
lines(density(y[g==2]), col = 2)
out <- sample_CAM(nrep = 500, y = y, group = g,
                 nclus_start = 2,
                 maxL = 20, maxK = 20)
out
```



sample\_fiSAN

*Sample fiSAN***Description**

sample\_fiSAN is used to perform posterior inference under the finite-infinite shared atoms nested (fiSAN) model with Gaussian likelihood. The model uses a Dirichlet process mixture prior at the distributional level, and Dirichlet mixture with an unknown number of components (following the specification of Frühwirth-Schnatter et al., 2021) at the observational level. The algorithm for the nonparametric component is based on the slice sampler for DPM of Kalli, Griffin and Walker (2011).

**Usage**

```
sample_fiSAN(nrep, y, group,
             maxK = 50, maxL = 50,
             m0 = 0, tau0 = 0.1, lambda0 = 3, gamma0 = 2,
             hyp_alpha1 = 1, hyp_alpha2 = 1,
             hyp_beta1 = 6, hyp_beta2 = 3,
             eps_beta = NULL,
             alpha = NULL, beta = NULL,
             warmstart = TRUE, nclus_start = NULL,
             mu_start = NULL, sigma2_start = NULL,
             M_start = NULL, S_start = NULL,
             alpha_start = NULL, beta_start = NULL,
             progress = TRUE, seed = NULL)
```

**Arguments**

nrep	Number of MCMC iterations.
y	Vector of observations.
group	Vector of the same length of y indicating the group membership (numeric).
maxK	Maximum number of distributional clusters $K$ (default = 50).
maxL	Maximum number of observational clusters $L$ (default = 50).
m0, tau0, lambda0, gamma0	Hyperparameters on $(\mu, \sigma^2) \sim NIG(m_0, \tau_0, \lambda_0, \gamma_0)$ . Default is (0, 0.1, 3, 2).
hyp_alpha1, hyp_alpha2	If a random $\alpha$ is used, (hyp_alpha1, hyp_alpha2) specify the hyperparameters (default = (1,1)). The prior is $\alpha \sim \text{Gamma}(\text{hyp\_alpha1}, \text{hyp\_alpha2})$ .
hyp_beta1, hyp_beta2, eps_beta	If a random $\beta$ is used, (hyp_beta1, hyp_beta2) specifies the hyperparameter (default = (6,3)). The prior is $\beta \sim \text{Gamma}(\text{hyp\_beta1}, \text{hyp\_beta2})$ . In this case, eps_beta is the tuning parameter of the MH step.
alpha	Distributional DP parameter if fixed (optional). The distribution is $\pi \sim \text{GEM}(\alpha)$ .

beta	Observational Dirichlet parameter if fixed (optional). The distribution is Dirichlet( $\text{rep}(\text{beta}/\text{maxL}, \text{maxL})$ ).
warmstart, nclus_start	Initialization of the observational clustering. warmstart is logical parameter (default = TRUE) of whether a kmeans clustering should be used to initialize the chains. An initial guess of the number of observational clusters can be passed via the nclus_start parameter (optional)
mu_start, sigma2_start, M_start, S_start, alpha_start, beta_start	Starting points of the MCMC chains (optional). Default is nclus_start = $\min(c(\text{maxL}, 30))$ . mu_start, sigma2_start are vectors of length maxL. M_start is a vector of observational cluster allocation of length N. S_start is a vector of observational cluster allocation of length J. alpha_start, alpha_start are numeric.
progress	show a progress bar? (logical, default TRUE).
seed	set a fixed seed.

## Details

### Data structure

The finite-infinite common atoms mixture model is used to perform inference in nested settings, where the data are organized into  $J$  groups. The data should be continuous observations  $(Y_1, \dots, Y_J)$ , where each  $Y_j = (y_{1,j}, \dots, y_{n_j,j})$  contains the  $n_j$  observations from group  $j$ , for  $j = 1, \dots, J$ . The function takes as input the data as a numeric vector  $y$  in this concatenated form. Hence  $y$  should be a vector of length  $n_1 + \dots + n_J$ . The group parameter is a numeric vector of the same size as  $y$  indicating the group membership for each individual observation. Notice that with this specification the observations in the same group need not be contiguous as long as the correspondence between the variables  $y$  and group is maintained.

### Model

The data are modeled using a univariate Gaussian likelihood, where both the mean and the variance are observational-cluster-specific, i.e.,

$$y_{i,j} \mid M_{i,j} = l \sim N(\mu_l, \sigma_l^2)$$

where  $M_{i,j} \in \{1, \dots, L\}$  is the observational cluster indicator of observation  $i$  in group  $j$ . The prior on the model parameters is a Normal-Inverse-Gamma distribution  $(\mu_l, \sigma_l^2) \sim \text{NIG}(m_0, \tau_0, \lambda_0, \gamma_0)$ , i.e.,  $\mu_l \mid \sigma_l^2 \sim N(m_0, \sigma_l^2/\tau_0)$ ,  $1/\sigma_l^2 \sim \text{Gamma}(\lambda_0, \gamma_0)$  (shape, rate).

### Clustering

The model performs a clustering of both observations and groups. The clustering of groups (distributional clustering) is provided by the allocation variables  $S_j \in \{1, 2, \dots\}$ , with

$$\Pr(S_j = k \mid \dots) = \pi_k \quad \text{for } k = 1, 2, \dots$$

The distribution of the probabilities is  $\{\pi_k\}_{k=1}^{\infty} \sim \text{GEM}(\alpha)$ , where GEM is the Griffiths-Engen-McCloskey distribution of parameter  $\alpha$ , which characterizes the stick-breaking construction of the DP (Sethuraman, 1994).

The clustering of observations (observational clustering) is provided by the allocation variables  $M_{i,j} \in \{1, \dots, L\}$ , with

$$\Pr(M_{i,j} = l \mid S_j = k, \dots) = \omega_{l,k} \quad \text{for } k = 1, 2, \dots; l = 1, \dots, L.$$

The distribution of the probabilities is  $(\omega_{1,k}, \dots, \omega_{L,k}) \sim \text{Dirichlet}_L(\beta/L, \dots, \beta/L)$  for all  $k = 1, 2, \dots$ . Moreover, the dimension  $L$  is random (see Frühwirth-Schnatter et al., 2021).

### Value

sample\_fiSAN returns four objects:

- `model`: name of the fitted model.
- `params`: list containing the data and the parameters used in the simulation. Details below.
- `sim`: list containing the simulated values (MCMC chains). Details below.
- `time`: total computation time.

**Data and parameters:** `params` is a list with the following components:

`nrep` Number of MCMC iterations.

`y, group` Data and group vectors.

`maxK, maxL` Maximum number of distributional and observational clusters.

`m0, tau0, lambda0, gamma0` Model hyperparameters.

`(hyp_alpha1, hyp_alpha2) or alpha` Either the hyperparameters on  $\alpha$  (if  $\alpha$  random), or the value for  $\alpha$  (if fixed).

`(hyp_beta1, hyp_beta2, eps_beta) or beta` Either the hyperparameters on  $\beta$  and MH step size (if  $\beta$  random), or the value for  $\beta$  (if fixed).

**Simulated values:** `sim` is a list with the following components:

`mu` Matrix of size  $(nrep, maxL)$ . Each row is a posterior sample of the mean parameter for each observational cluster  $(\mu_1, \dots, \mu_L)$ .

`sigma2` Matrix of size  $(nrep, maxL)$ . Each row is a posterior sample of the variance parameter for each observational cluster  $(\sigma_1^2, \dots, \sigma_L^2)$ .

`obs_cluster` Matrix of size  $(nrep, n)$ , with  $n = \text{length}(y)$ . Each row is a posterior sample of the observational cluster allocation variables  $(M_{1,1}, \dots, M_{n,J})$ .

`distr_cluster` Matrix of size  $(nrep, J)$ , with  $J = \text{length}(\text{unique}(\text{group}))$ . Each row is a posterior sample of the distributional cluster allocation variables  $(S_1, \dots, S_J)$ .

`pi` Matrix of size  $(nrep, maxK)$ . Each row is a posterior sample of the distributional cluster probabilities  $(\pi_1, \dots, \pi_{maxK})$ .

`omega` 3-d array of size  $(maxL, maxK, nrep)$ . Each slice is a posterior sample of the observational cluster probabilities. In each slice, each column  $k$  is a vector (of length  $maxL$ ) observational cluster probabilities  $(\omega_{1,k}, \dots, \omega_{L,k})$  for distributional cluster  $k$ .

`alpha` Vector of length `nrep` of posterior samples of the parameter  $\alpha$ .

`beta` Vector of length `nrep` of posterior samples of the parameter  $\beta$ .

`maxK` Vector of length `nrep` of the number of distributional DP components used by the slice sampler.

`L_plus` Vector of length `nrep` of posterior samples of the number of observational clusters.

`L` Vector of length `nrep` of posterior samples of the observational Dirichlet dimension.

## References

- Frühwirth-Schnatter, S., Malsiner-Walli, G. and Grün, B. (2021). Generalized mixtures of finite mixtures and telescoping sampling. *Bayesian Analysis*, 16(4), 1279–1307. <doi:10.1214/21-BA1294>
- Kalli, M., Griffin, J.E., and Walker, S.G. (2011). Slice Sampling Mixture Models, *Statistics and Computing*, 21, 93–105. <doi:10.1007/s11222-009-9150-y>
- Sethuraman, A.J. (1994). A Constructive Definition of Dirichlet Priors, *Statistica Sinica*, 4, 639–650.

## Examples

```
set.seed(123)
y <- c(rnorm(40,0,0.3), rnorm(20,5,0.3))
g <- c(rep(1,30), rep(2, 30))
plot(density(y[g==1]), xlim = c(-5,10))
lines(density(y[g==2]), col = 2)
out <- sample_fiSAN(nrep = 500, y = y, group = g,
                   nclus_start = 2,
                   maxK = 20, maxL = 20,
                   beta = 1)
out
```

---

sample\_fiSAN\_sparsemix

*Sample fiSAN with sparse mixtures*

---

## Description

sample\_fiSAN\_sparsemix is used to perform posterior inference under the finite-infinite shared atoms nested (fiSAN) model with Gaussian likelihood. The model uses a Dirichlet process mixture prior at the distributional level, and a sparse (overfitted) Dirichlet mixture (Malsiner-Walli et al., 2016) at the observational level. The algorithm for the nonparametric component is based on the slice sampler for DPM of Kalli, Griffin and Walker (2011).

## Usage

```
sample_fiSAN_sparsemix(nrep, y, group,
                      maxK = 50, maxL = 50,
                      m0 = 0, tau0 = 0.1, lambda0 = 3, gamma0 = 2,
                      hyp_alpha1 = 1, hyp_alpha2 = 1,
                      hyp_beta = 10,
                      eps_beta = NULL,
                      alpha = NULL, beta = NULL,
                      warmstart = TRUE, nclus_start = NULL,
                      mu_start = NULL, sigma2_start = NULL,
                      M_start = NULL, S_start = NULL,
                      alpha_start = NULL, beta_start = NULL,
                      progress = TRUE, seed = NULL)
```

**Arguments**

nrep	Number of MCMC iterations.
y	Vector of observations.
group	Vector of the same length of y indicating the group membership (numeric).
maxK	Maximum number of distributional clusters $K$ (default = 50).
maxL	Maximum number of observational clusters $L$ (default = 50).
$m_0$ , $\tau_0$ , $\lambda_0$ , $\gamma_0$	Hyperparameters on $(\mu, \sigma^2) \sim NIG(m_0, \tau_0, \lambda_0, \gamma_0)$ . Default is (0, 0.1, 3, 2).
hyp_alpha1, hyp_alpha2	If a random $\alpha$ is used, (hyp_alpha1, hyp_alpha2) specify the hyperparameters (default = (1,1)). The prior is $\alpha \sim \text{Gamma}(\text{hyp\_alpha1}, \text{hyp\_alpha2})$ .
hyp_beta, eps_beta	If a random $\beta$ is used, hyp_beta specifies the hyperparameter (default = 10). The prior is $\beta \sim \text{Gamma}(\text{hyp\_beta}, \text{hyp\_beta} * \text{maxL})$ , following Malsiner-Walli et al. (2016). In this case, eps_beta is the tuning parameter of the MH step.
alpha	Distributional DP parameter if fixed (optional). The distribution is $\pi \sim GEM(\alpha)$ .
beta	Observational Dirichlet parameter if fixed (optional). The distribution is $\text{Dirichlet}(\text{rep}(\text{beta}, \text{maxL}))$ . Notice that beta should be small to ensure sparsity (e.g. beta = 0.01)
warmstart, nclus_start	Initialization of the observational clustering. warmstart is logical parameter (default = TRUE) of whether a kmeans clustering should be used to initialize the chains. An initial guess of the number of observational clusters can be passed via the nclus_start parameter (optional)
mu_start, sigma2_start, M_start, S_start, alpha_start, beta_start	Starting points of the MCMC chains (optional). Default is nclus_start = $\min(c(\text{maxL}, 30))$ . mu_start, sigma2_start are vectors of length maxL. M_start is a vector of observational cluster allocation of length N. S_start is a vector of observational cluster allocation of length J. alpha_start, alpha_start are numeric.
progress	show a progress bar? (logical, default TRUE).
seed	set a fixed seed.

**Details****Data structure**

The finite-infinite common atoms mixture model is used to perform inference in nested settings, where the data are organized into  $J$  groups. The data should be continuous observations  $(Y_1, \dots, Y_J)$ , where each  $Y_j = (y_{1,j}, \dots, y_{n_j,j})$  contains the  $n_j$  observations from group  $j$ , for  $j = 1, \dots, J$ . The function takes as input the data as a numeric vector  $y$  in this concatenated form. Hence  $y$  should be a vector of length  $n_1 + \dots + n_J$ . The group parameter is a numeric vector of the same size as  $y$  indicating the group membership for each individual observation. Notice that with this specification the observations in the same group need not be contiguous as long as the correspondence between the variables  $y$  and  $group$  is maintained.

**Model**

The data are modeled using a univariate Gaussian likelihood, where both the mean and the variance are observational-cluster-specific, i.e.,

$$y_{i,j} \mid M_{i,j} = l \sim N(\mu_l, \sigma_l^2)$$

where  $M_{i,j} \in \{1, \dots, L\}$  is the observational cluster indicator of observation  $i$  in group  $j$ . The prior on the model parameters is a Normal-Inverse-Gamma distribution  $(\mu_l, \sigma_l^2) \sim NIG(m_0, \tau_0, \lambda_0, \gamma_0)$ , i.e.,  $\mu_l \mid \sigma_l^2 \sim N(m_0, \sigma_l^2/\tau_0)$ ,  $1/\sigma_l^2 \sim \text{Gamma}(\lambda_0, \gamma_0)$  (shape, rate).

**Clustering**

The model performs a clustering of both observations and groups. The clustering of groups (distributional clustering) is provided by the allocation variables  $S_j \in \{1, 2, \dots\}$ , with

$$Pr(S_j = k \mid \dots) = \pi_k \quad \text{for } k = 1, 2, \dots$$

The distribution of the probabilities is  $\{\pi_k\}_{k=1}^{\infty} \sim GEM(\alpha)$ , where GEM is the Griffiths-Engen-McCloskey distribution of parameter  $\alpha$ , which characterizes the stick-breaking construction of the DP (Sethuraman, 1994).

The clustering of observations (observational clustering) is provided by the allocation variables  $M_{i,j} \in \{1, \dots, L\}$ , with

$$Pr(M_{i,j} = l \mid S_j = k, \dots) = \omega_{l,k} \quad \text{for } k = 1, 2, \dots; l = 1, \dots, L.$$

The distribution of the probabilities is  $(\omega_{1,k}, \dots, \omega_{L,k}) \sim \text{Dirichlet}_L(\beta, \dots, \beta)$  for all  $k = 1, 2, \dots$ .

**Value**

sample\_fiSAN\_sparsemix returns four objects:

- model: name of the fitted model.
- params: list containing the data and the parameters used in the simulation. Details below.
- sim: list containing the simulated values (MCMC chains). Details below.
- time: total computation time.

**Data and parameters:** params is a list with the following components:

nrep Number of MCMC iterations.

y, group Data and group vectors.

maxK, maxL Maximum number of distributional and observational clusters.

m0, tau0, lambda0, gamma0 Model hyperparameters.

(hyp\_alpha1, hyp\_alpha2) **or** alpha Either the hyperparameters on  $\alpha$  (if  $\alpha$  random), or the value for  $\alpha$  (if fixed).

(hyp\_beta, eps\_beta) **or** beta Either the hyperparameter on  $\beta$  and MH step size (if  $\beta$  random), or the value for  $\beta$  (if fixed).

**Simulated values:** sim is a list with the following components:

- mu** Matrix of size (nrep, maxL). Each row is a posterior sample of the mean parameter for each observational cluster  $(\mu_1, \dots, \mu_L)$ .
- sigma2** Matrix of size (nrep, maxL). Each row is a posterior sample of the variance parameter for each observational cluster  $(\sigma_1^2, \dots, \sigma_L^2)$ .
- obs\_cluster** Matrix of size (nrep, n), with  $n = \text{length}(y)$ . Each row is a posterior sample of the observational cluster allocation variables  $(M_{1,1}, \dots, M_{n,J})$ .
- distr\_cluster** Matrix of size (nrep, J), with  $J = \text{length}(\text{unique}(\text{group}))$ . Each row is a posterior sample of the distributional cluster allocation variables  $(S_1, \dots, S_J)$ .
- pi** Matrix of size (nrep, maxK). Each row is a posterior sample of the distributional cluster probabilities  $(\pi_1, \dots, \pi_{\text{maxK}})$ .
- omega** 3-d array of size (maxL, maxK, nrep). Each slice is a posterior sample of the observational cluster probabilities. In each slice, each column  $k$  is a vector (of length maxL) observational cluster probabilities  $(\omega_{1,k}, \dots, \omega_{L,k})$  for distributional cluster  $k$ .
- alpha** Vector of length nrep of posterior samples of the parameter  $\alpha$ .
- beta** Vector of length nrep of posterior samples of the parameter  $\beta$ .
- maxK** Vector of length nrep of the number of distributional DP components used by the slice sampler.

## References

- Kalli, M., Griffin, J.E., and Walker, S.G. (2011). Slice Sampling Mixture Models, *Statistics and Computing*, 21, 93–105. <doi:10.1007/s11222-009-9150-y>
- Malsiner-Walli, G., Frühwirth-Schnatter, S. and Grün, B. (2016). Model-based clustering based on sparse finite Gaussian mixtures. *Statistics and Computing* 26, 303–324. <doi:10.1007/s11222-014-9500-2>
- Sethuraman, A.J. (1994). A Constructive Definition of Dirichlet Priors, *Statistica Sinica*, 4, 639–650.

## Examples

```
set.seed(123)
y <- c(rnorm(40,0,0.3), rnorm(20,5,0.3))
g <- c(rep(1,30), rep(2, 30))
plot(density(y[g==1]), xlim = c(-5,10))
lines(density(y[g==2]), col = 2)
out <- sample_fiSAN_sparsemix(nrep = 500, y = y, group = g,
                             nclus_start = 2,
                             maxK = 20, maxL = 20,
                             beta = 0.01)
out
```

sample\_fSAN

*Sample fSAN***Description**

sample\_fSAN is used to perform posterior inference under the finite shared atoms nested (fSAN) model with Gaussian likelihood (originally proposed in D’Angelo et al., 2023). The model uses Dirichlet mixtures with an unknown number of components (following the specification of Frühwirth-Schnatter et al., 2021) at both the observational and distributional level.

**Usage**

```
sample_fSAN(nrep, y, group,
            maxK = 50, maxL = 50,
            m0 = 0, tau0 = 0.1, lambda0 = 3, gamma0 = 2,
            hyp_alpha1 = 6, hyp_alpha2 = 3,
            hyp_beta1 = 6, hyp_beta2 = 3,
            eps_alpha = NULL, eps_beta = NULL,
            alpha = NULL, beta = NULL,
            warmstart = TRUE, nclus_start = NULL,
            mu_start = NULL, sigma2_start = NULL,
            M_start = NULL, S_start = NULL,
            alpha_start = NULL, beta_start = NULL,
            progress = TRUE, seed = NULL)
```

**Arguments**

nrep	Number of MCMC iterations.
y	Vector of observations.
group	Vector of the same length of y indicating the group membership (numeric).
maxK	Maximum number of distributional clusters $K$ (default = 50).
maxL	Maximum number of observational clusters $L$ (default = 50).
$m_0$ , $\tau_0$ , $\lambda_0$ , $\gamma_0$	Hyperparameters on $(\mu, \sigma^2) \sim NIG(m_0, \tau_0, \lambda_0, \gamma_0)$ . Default is (0, 0.1, 3, 2).
hyp_alpha1, hyp_alpha2, eps_alpha	If a random $\alpha$ is used, (hyp_alpha1, hyp_alpha2) specify the hyperparameters (default = (6,3)). The prior is $\alpha \sim Gamma(\text{hyp\_alpha1}, \text{hyp\_alpha2})$ . In this case, eps_alpha is the tuning parameter of the MH step.
hyp_beta1, hyp_beta2, eps_beta	If a random $\beta$ is used, (hyp_beta1, hyp_beta2) specifies the hyperparameter (default = (6,3)). The prior is $\beta \sim Gamma(\text{hyp\_beta1}, \text{hyp\_beta2})$ . In this case, eps_beta is the tuning parameter of the MH step.
alpha	Distributional Dirichlet parameter if fixed (optional). The distribution is $Dirichlet(\text{rep}(\text{alpha}/\text{maxK}, \text{maxK}))$ .



beta	Observational Dirichlet parameter if fixed (optional). The distribution is Dirichlet( $\text{rep}(\text{beta}/\text{maxL}, \text{maxL})$ ).
warmstart, nclus_start	Initialization of the observational clustering. warmstart is logical parameter (default = TRUE) of whether a kmeans clustering should be used to initialize the chains. An initial guess of the number of observational clusters can be passed via the nclus_start parameter (optional). Default is nclus_start = $\min(c(\text{maxL}, 30))$ .
mu_start, sigma2_start, M_start, S_start, alpha_start, beta_start	Starting points of the MCMC chains (optional). mu_start, sigma2_start are vectors of length maxL. M_start is a vector of observational cluster allocation of length N. S_start is a vector of observational cluster allocation of length J. alpha_start, beta_start are numeric.
progress	show a progress bar? (logical, default TRUE.)
seed	set a fixed seed.

## Details

### Data structure

The finite common atoms mixture model is used to perform inference in nested settings, where the data are organized into  $J$  groups. The data should be continuous observations  $(Y_1, \dots, Y_J)$ , where each  $Y_j = (y_{1,j}, \dots, y_{n_j,j})$  contains the  $n_j$  observations from group  $j$ , for  $j = 1, \dots, J$ . The function takes as input the data as a numeric vector  $y$  in this concatenated form. Hence  $y$  should be a vector of length  $n_1 + \dots + n_J$ . The group parameter is a numeric vector of the same size as  $y$  indicating the group membership for each individual observation. Notice that with this specification the observations in the same group need not be contiguous as long as the correspondence between the variables  $y$  and group is maintained.

### Model

The data are modeled using a univariate Gaussian likelihood, where both the mean and the variance are observational-cluster-specific, i.e.,

$$y_{i,j} \mid M_{i,j} = l \sim N(\mu_l, \sigma_l^2)$$

where  $M_{i,j} \in \{1, \dots, L\}$  is the observational cluster indicator of observation  $i$  in group  $j$ . The prior on the model parameters is a Normal-Inverse-Gamma distribution  $(\mu_l, \sigma_l^2) \sim \text{NIG}(m_0, \tau_0, \lambda_0, \gamma_0)$ , i.e.,  $\mu_l \mid \sigma_l^2 \sim N(m_0, \sigma_l^2/\tau_0)$ ,  $1/\sigma_l^2 \sim \text{Gamma}(\lambda_0, \gamma_0)$  (shape, rate).

### Clustering

The model performs a clustering of both observations and groups. The clustering of groups (distributional clustering) is provided by the allocation variables  $S_j \in \{1, \dots, K\}$ , with

$$\Pr(S_j = k \mid \dots) = \pi_k \quad \text{for } k = 1, \dots, K.$$

The distribution of the probabilities is  $(\pi_1, \dots, \pi_K) \sim \text{Dirichlet}_K(\alpha/K, \dots, \alpha/K)$ . Moreover, the dimension  $K$  is random (see Frühwirth-Schnatter et al., 2021).

The clustering of observations (observational clustering) is provided by the allocation variables  $M_{i,j} \in \{1, \dots, L\}$ , with

$$\Pr(M_{i,j} = l \mid S_j = k, \dots) = \omega_{l,k} \quad \text{for } k = 1, \dots, K; l = 1, \dots, L.$$

The distribution of the probabilities is  $(\omega_{1,k}, \dots, \omega_{L,k}) \sim \text{Dirichlet}_L(\beta/L, \dots, \beta/L)$  for all  $k = 1, \dots, K$ . Moreover, the dimension  $L$  is random (see Frühwirth-Schnatter et al., 2021).

### Value

sample\_fSAN returns four objects:

- model: name of the fitted model.
- params: list containing the data and the parameters used in the simulation. Details below.
- sim: list containing the simulated values (MCMC chains). Details below.
- time: total computation time.

**Data and parameters:** params is a list with the following components:

nrep Number of MCMC iterations.

y, group Data and group vectors.

maxK, maxL Maximum number of distributional and observational clusters.

m0, tau0, lambda0, gamma0 Model hyperparameters.

(hyp\_alpha1, hyp\_alpha2, eps\_alpha) **or** alpha Either the hyperparameters on  $\alpha$  and MH step size (if  $\alpha$  random), or the value for  $\alpha$  (if fixed).

(hyp\_beta1, hyp\_beta2, eps\_beta) **or** beta Either the hyperparameters on  $\beta$  and MH step size (if  $\beta$  random), or the value for  $\beta$  (if fixed).

**Simulated values:** sim is a list with the following components:

mu Matrix of size (nrep, maxL). Each row is a posterior sample of the mean parameter for each observational cluster  $(\mu_1, \dots, \mu_L)$ .

sigma2 Matrix of size (nrep, maxL). Each row is a posterior sample of the variance parameter for each observational cluster  $(\sigma_1^2, \dots, \sigma_L^2)$ .

obs\_cluster Matrix of size (nrep, n), with  $n = \text{length}(y)$ . Each row is a posterior sample of the observational cluster allocation variables  $(M_{1,1}, \dots, M_{n,J})$ .

distr\_cluster Matrix of size (nrep, J), with  $J = \text{length}(\text{unique}(\text{group}))$ . Each row is a posterior sample of the distributional cluster allocation variables  $(S_1, \dots, S_J)$ .

pi Matrix of size (nrep, maxK). Each row is a posterior sample of the distributional cluster probabilities  $(\pi_1, \dots, \pi_K)$ .

omega 3-d array of size (maxL, maxK, nrep). Each slice is a posterior sample of the observational cluster probabilities. In each slice, each column  $k$  is a vector (of length maxL) observational cluster probabilities  $(\omega_{1,k}, \dots, \omega_{L,k})$  for distributional cluster  $k$ .

alpha Vector of length nrep of posterior samples of the parameter  $\alpha$ .

beta Vector of length nrep of posterior samples of the parameter  $\beta$ .

K\_plus Vector of length nrep of posterior samples of the number of distributional clusters.

L\_plus Vector of length nrep of posterior samples of the number of observational clusters.

K Vector of length nrep of posterior samples of the distributional Dirichlet dimension.

L Vector of length nrep of posterior samples of the observational Dirichlet dimension.

## References

D'Angelo, L., Canale, A., Yu, Z., and Guindani, M. (2023). Bayesian nonparametric analysis for the detection of spikes in noisy calcium imaging data. *Biometrics*, 79(2), 1370–1382. <doi:10.1111/biom.13626>

Frühwirth-Schnatter, S., Malsiner-Walli, G. and Grün, B. (2021). Generalized mixtures of finite mixtures and telescoping sampling. *Bayesian Analysis*, 16(4), 1279–1307. <doi:10.1214/21-BA1294>

## Examples

```
set.seed(123)
y <- c(rnorm(40,0,0.3), rnorm(20,5,0.3))
g <- c(rep(1,30), rep(2, 30))
plot(density(y[g==1]), xlim = c(-5,10))
lines(density(y[g==2]), col = 2)
out <- sample_fSAN(nrep = 500, y = y, group = g,
                  nclus_start = 2,
                  maxK = 20, maxL = 20,
                  alpha = 1, beta = 1)
out
```

---

sample\_fSAN\_sparsemix *Sample fSAN with sparse mixtures*

---

## Description

sample\_fSAN\_sparsemix is used to perform posterior inference under the finite shared atoms nested (fSAN) model with Gaussian likelihood (originally proposed in D'Angelo et al., 2023). The model uses overfitted (sparse) Dirichlet mixtures (Malsiner-Walli et al., 2016) at both the observational and distributional level.

## Usage

```
sample_fSAN_sparsemix(nrep, y, group,
                    maxK = 50, maxL = 50,
                    m0 = 0, tau0 = 0.1, lambda0 = 3, gamma0 = 2,
                    hyp_alpha = 10, hyp_beta = 10,
                    eps_alpha = NULL, eps_beta = NULL,
                    alpha = NULL, beta = NULL,
                    warmstart = TRUE, nclus_start = NULL,
                    mu_start = NULL, sigma2_start = NULL,
                    M_start = NULL, S_start = NULL,
                    alpha_start = NULL, beta_start = NULL,
                    progress = TRUE, seed = NULL)
```

**Arguments**

nrep	Number of MCMC iterations.
y	Vector of observations.
group	Vector of the same length of y indicating the group membership (numeric).
maxK	Maximum number of distributional clusters $K$ (default = 50).
maxL	Maximum number of observational clusters $L$ (default = 50).
$m_0$ , $\tau_0$ , $\lambda_0$ , $\gamma_0$	Hyperparameters on $(\mu, \sigma^2) \sim NIG(m_0, \tau_0, \lambda_0, \gamma_0)$ . Default is (0, 0.1, 3, 2).
hyp_alpha, eps_alpha	If a random $\alpha$ is used, hyp_alpha specifies the hyperparameter (default = 10). The prior is $\alpha \sim \text{Gamma}(\text{hyp\_alpha}, \text{hyp\_alpha} * \text{maxK})$ , following Malsiner-Walli et al. (2016). In this case, eps_alpha is the tuning parameter of the MH step.
hyp_beta, eps_beta	If a random $\beta$ is used, hyp_beta specifies the hyperparameter (default = 10). The prior is $\beta \sim \text{Gamma}(\text{hyp\_beta}, \text{hyp\_beta} * \text{maxL})$ , following Malsiner-Walli et al. (2016). In this case, eps_beta is the tuning parameter of the MH step.
alpha	Distributional Dirichlet parameter if fixed (optional). The distribution is Dirichlet( rep(alpha, maxK) ).
beta	Observational Dirichlet parameter if fixed (optional). The distribution is Dirichlet( rep(beta, maxL) ).
warmstart, nclus_start	Initialization of the observational clustering. warmstart is logical parameter (default = TRUE) of whether a kmeans clustering should be used to initialize the chains. An initial guess of the number of observational clusters can be passed via the nclus_start parameter (optional)
mu_start, sigma2_start, M_start, S_start, alpha_start, beta_start	Starting points of the MCMC chains (optional). Default is nclus_start = min(c(maxL, 30)). mu_start, sigma2_start are vectors of length maxL. M_start is a vector of observational cluster allocation of length N. S_start is a vector of observational cluster allocation of length J. alpha_start, alpha_start are numeric.
progress	show a progress bar? (logical, default TRUE).
seed	set a fixed seed.

**Details****Data structure**

The overfitted mixture common atoms model is used to perform inference in nested settings, where the data are organized into  $J$  groups. The data should be continuous observations  $(Y_1, \dots, Y_J)$ , where each  $Y_j = (y_{1,j}, \dots, y_{n_j,j})$  contains the  $n_j$  observations from group  $j$ , for  $j = 1, \dots, J$ . The function takes as input the data as a numeric vector  $y$  in this concatenated form. Hence  $y$  should be a vector of length  $n_1 + \dots + n_J$ . The group parameter is a numeric vector of the same size as  $y$  indicating the group membership for each individual observation. Notice that with this specification

the observations in the same group need not be contiguous as long as the correspondence between the variables  $y$  and group is maintained.

### Model

The data are modeled using a univariate Gaussian likelihood, where both the mean and the variance are observational-cluster-specific, i.e.,

$$y_{i,j} \mid M_{i,j} = l \sim N(\mu_l, \sigma_l^2)$$

where  $M_{i,j} \in \{1, \dots, L\}$  is the observational cluster indicator of observation  $i$  in group  $j$ . The prior on the model parameters is a Normal-Inverse-Gamma distribution  $(\mu_l, \sigma_l^2) \sim NIG(m_0, \tau_0, \lambda_0, \gamma_0)$ , i.e.,  $\mu_l \mid \sigma_l^2 \sim N(m_0, \sigma_l^2/\tau_0)$ ,  $1/\sigma_l^2 \sim \text{Gamma}(\lambda_0, \gamma_0)$  (shape, rate).

### Clustering

The model performs a clustering of both observations and groups. The clustering of groups (distributional clustering) is provided by the allocation variables  $S_j \in \{1, \dots, K\}$ , with

$$\Pr(S_j = k \mid \dots) = \pi_k \quad \text{for } k = 1, \dots, K.$$

The distribution of the probabilities is  $(\pi_1, \dots, \pi_K) \sim \text{Dirichlet}_K(\alpha, \dots, \alpha)$ .

The clustering of observations (observational clustering) is provided by the allocation variables  $M_{i,j} \in \{1, \dots, L\}$ , with

$$\Pr(M_{i,j} = l \mid S_j = k, \dots) = \omega_{l,k} \quad \text{for } k = 1, \dots, K; l = 1, \dots, L.$$

The distribution of the probabilities is  $(\omega_{1,k}, \dots, \omega_{L,k}) \sim \text{Dirichlet}_L(\beta, \dots, \beta)$  for all  $k = 1, \dots, K$ .

### Value

sample\_fSAN\_sparsemix returns four objects:

- model: name of the fitted model.
- params: list containing the data and the parameters used in the simulation. Details below.
- sim: list containing the simulated values (MCMC chains). Details below.
- time: total computation time.

**Data and parameters:** params is a list with the following components:

nrep Number of MCMC iterations.

y, group Data and group vectors.

maxK, maxL Maximum number of distributional and observational clusters.

m0, tau0, lambda0, gamma0 Model hyperparameters.

(hyp\_alpha, eps\_alpha) **or** alpha Either the hyperparameter on  $\alpha$  and MH step size (if  $\alpha$  random), or the value for  $\alpha$  (if fixed).

(hyp\_beta, eps\_beta) **or** beta Either the hyperparameter on  $\beta$  and MH step size (if  $\beta$  random), or the value for  $\beta$  (if fixed).

**Simulated values:** sim is a list with the following components:

- mu** Matrix of size (nrep, maxL). Each row is a posterior sample of the mean parameter for each observational cluster  $(\mu_1, \dots, \mu_L)$ .
- sigma2** Matrix of size (nrep, maxL). Each row is a posterior sample of the variance parameter for each observational cluster  $(\sigma_1^2, \dots, \sigma_L^2)$ .
- obs\_cluster** Matrix of size (nrep, n), with  $n = \text{length}(y)$ . Each row is a posterior sample of the observational cluster allocation variables  $(M_{1,1}, \dots, M_{n,J})$ .
- distr\_cluster** Matrix of size (nrep, J), with  $J = \text{length}(\text{unique}(\text{group}))$ . Each row is a posterior sample of the distributional cluster allocation variables  $(S_1, \dots, S_J)$ .
- pi** Matrix of size (nrep, maxK). Each row is a posterior sample of the distributional cluster probabilities  $(\pi_1, \dots, \pi_K)$ .
- omega** 3-d array of size (maxL, maxK, nrep). Each slice is a posterior sample of the observational cluster probabilities. In each slice, each column  $k$  is a vector (of length maxL) observational cluster probabilities  $(\omega_{1,k}, \dots, \omega_{L,k})$  for distributional cluster  $k$ .
- alpha** Vector of length nrep of posterior samples of the parameter  $\alpha$ .
- beta** Vector of length nrep of posterior samples of the parameter  $\beta$ .

## References

- D'Angelo, L., Canale, A., Yu, Z., and Guindani, M. (2023). Bayesian nonparametric analysis for the detection of spikes in noisy calcium imaging data. *Biometrics*, 79(2), 1370–1382. <doi:10.1111/biom.13626>
- Malsiner-Walli, G., Frühwirth-Schnatter, S. and Grün, B. (2016). Model-based clustering based on sparse finite Gaussian mixtures. *Statistics and Computing* 26, 303–324. <doi:10.1007/s11222-014-9500-2>

## Examples

```
set.seed(123)
y <- c(rnorm(40,0,0.3), rnorm(20,5,0.3))
g <- c(rep(1,30), rep(2, 30))
plot(density(y[g==1]), xlim = c(-5,10))
lines(density(y[g==2]), col = 2)
out <- sample_fSAN_sparsemix(nrep = 500, y = y, group = g,
                             nclus_start = 2,
                             maxK = 20, maxL = 20,
                             alpha = 0.01, beta = 0.01)

out
```

---

traceplot

*Traceplot: plot MCMC chains*

---

## Description

Check the convergence of the MCMC through visual inspection of the chains.

**Usage**

```
traceplot(object, params,
          show_density = TRUE,
          show_burnin = TRUE,
          length_burnin = NULL,
          show_convergence = TRUE,
          trunc_plot = 10)
```

**Arguments**

object	object of class SANmcmc (the result of a call to <a href="#">sample_fiSAN</a> , <a href="#">sample_fiSAN_sparsemix</a> , <a href="#">sample_fSAN</a> , <a href="#">sample_fSAN_sparsemix</a> , or <a href="#">sample_CAM</a> ).
params	vector of strings with the names of the parameters to check.
show_density	logical (default TRUE). Whether a kernel estimate of the density should be plotted. The burn-in is always discarded.
show_burnin	logical (default TRUE). Whether the first part of the chains should be plotted in the traceplots.
length_burnin	if show_burnin = FALSE, the length of the burn-in to be discarded.
show_convergence	logical (default TRUE). Whether a superimposed red line of the cumulative mean should be plotted.
trunc_plot	integer (default = 10). For multidimensional parameters, the maximum number of components to be plotted.

**Value**

The function displays the traceplots of the MCMC algorithm.

**Note**

The function is not available for the observational weights  $\omega$ .

**Examples**

```
set.seed(123)
y <- c(rnorm(40,0,0.3), rnorm(20,5,0.3))
g <- c(rep(1,30), rep(2, 30))
out <- sample_fiSAN(nrep = 500, y = y, group = g,
                  nclus_start = 2,
                  maxK = 20, maxL = 20,
                  beta = 1)
traceplot(out, params = c("mu", "sigma2"), trunc_plot = 2)
```

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