

# Parameter estimation in a subcritical percolation model with colouring

Albert-Ludwigs-Universität Freiburg



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FREIBURG

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jointly with Felix Beck; arXiv:1604.08908 [math.ST]

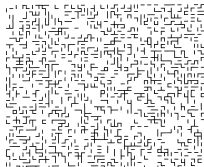
Centre for Biological Systems Analysis (ZBSA)

June 2016

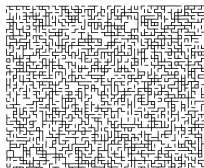
# Bond percolation

4 What is Percolation?

[1.2]



(a)  $p = 0.25$

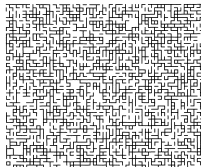


(b)  $p = 0.49$

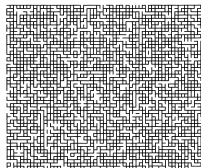
Figure 1.2. Realizations of bond percolation on a  $50 \times 60$  section of the square lattice for four different values of  $p$ . The pictures have been created using the same sequence of pseudo-random numbers, with the result that each graph is a subgraph of the next. Realizers with good

[1.2]

Why Percolation? 5



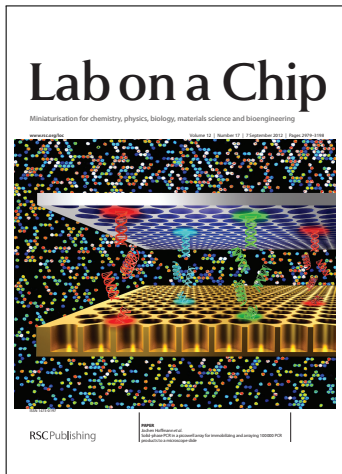
(c)  $p = 0.51$



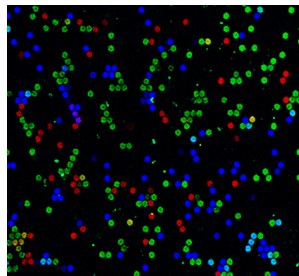
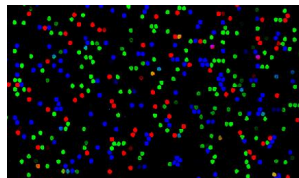
(d)  $p = 0.75$

eyesight may care to check that there exist open paths joining the left to the right side when  $p = 0.51$  but not when  $p = 0.49$ . The (random) value of  $p$  at which such paths appear for this realization is 0.5059. ...

# Cross-contamination rate estimation for digital PCR in lab-on-a-chip microfluidic devices



Hoffmann *et al.*; Lab on a Chip, 12, 3049–3054, 2012.



Rath; MSc thesis, Univ. of Freiburg, 2014.

## Contamination:

- (i) unidirectional (independent, directed edges  $\xi_{i \rightarrow j}$  and  $\xi_{j \rightarrow i}$  between any two adjacent vertices  $i \sim j$ ),  
or
- (ii) symmetric (undirected edges  $\xi_{ij}$ ).

## Open edges:

- (1) independent Bernoulli variables, or
- (2) locally correlated 0–1 random variables.

## Contamination:

- (A) confined to neighbours, or
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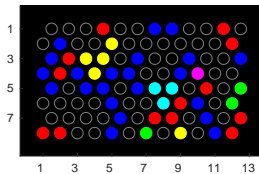
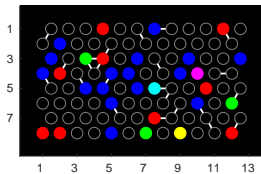
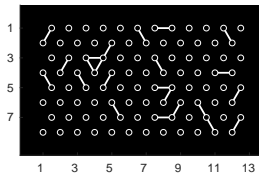
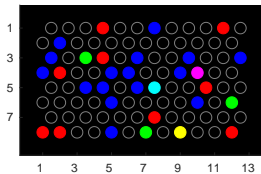
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# The process



- Process: index set  $I$  ( $n_I := |I|$ ), colours  $\ell \in \{1, 2, \dots, n_c\}$

$$Y_i^\ell := X_i^\ell \vee \bigvee_{j: j \leftrightarrow i} X_j^\ell, \quad Y_i^\ell \in \{0, 1\}$$

- Goal: estimate  $\theta = (\lambda^1, \dots, \lambda^{n_c}, \mu)$  from the data  $(Y_i^\ell)_{i \in I, \ell \in \{1, 2, \dots, n_c\}}$

- 1 Objective: the method of simulated moments (MSM) is strongly consistent.
- 2 We prove a strong law of large numbers (SLLN) with weakly dependent variables.
- 3 To upper bound dependence (i.e. correlations bw. vertices), we use the FKG and BK inequalities of percolation theory.



# The method of simulated moments (MSM)

- Data  $\mathcal{Y} = (\mathcal{Y}_i)_{i \in I}$  originates from a distribution which is parameterised by the unknown  $\theta_0 \in \Theta$  (the *true parameter value*).
- **Moments:**  $K$  is some  $n_m$ -dimensional function of the individual observations  $Y_i$ .  $k(\theta) := E_\theta[K(Y_i)]$
- **Identifiability:**  $E_{\theta_0}[K(\mathcal{Y}_i)] = k(\theta) \iff \theta = \theta_0$
- **MSM:**  $k(\theta)$  is not available in analytical form but there exists an unbiased estimator  $\tilde{k}(U_i^s, \theta)$ .  $(U_i^s)_{i \in I, s \in \{1, \dots, n_s\}}$  is some source of randomness, typically vectors of independent  $U[0, 1]$ .
- $\Omega \in \mathbb{R}^{n_m \times n_m}$  is a symmetric, positive definite matrix;  $\alpha(\eta) = \eta^T \Omega \eta$  a quadratic form. The MSM estimator is

$$\hat{\theta}_{n_s, n_I} := \arg \min_{\theta \in \Theta} \alpha \left( \frac{1}{n_I} \sum_{i=1}^{n_I} \left( K(\mathcal{Y}_i) - \frac{1}{n_s} \sum_{s=1}^{n_s} \tilde{k}(U_i^s, \theta) \right) \right)$$

# The method of simulated moments (MSM)



## Proposition

$\Omega \in \mathbb{R}^{n_m \times n_m}$  is symmetric, positive definite;  $\alpha(\eta) = \eta^T \Omega \eta$ .  
The MSM estimator is

$$\hat{\theta}_{n_s, n_l} := \arg \min_{\theta \in \Theta} \alpha \left( \frac{1}{n_l} \sum_{i=1}^{n_l} \left( K(\mathcal{Y}_i) - \frac{1}{n_s} \sum_{s=1}^{n_s} \tilde{k}(U_i^s, \theta) \right) \right).$$

If  $n_s$  is fixed and  $n_l$  tends to infinity, and the almost sure convergence guaranteed by the SLLN

$$\frac{1}{n_l} \sum_{i=1}^{n_l} \tilde{k}(U_i^s, \theta) \xrightarrow{n_l \rightarrow \infty} k(\theta)$$

is uniform in  $\theta \in \Theta$  for every  $s$ , then  $\hat{\theta}_{n_s, n_l}$  is strongly consistent (i.e.  $\hat{\theta}_{n_s, n_l}$  converges to  $\theta_0$  almost surely).

- Variables  $(Y_i^\ell)$  are neither identically distributed (*boundary!*) nor independent. — SLLN is not a given.
- Moments we use:  $Y_i^\ell, Y_i^\ell Y_j^\ell$  for  $i \sim j$
- Identifiability:  
If  $(\lambda^1, \dots, \lambda^{n_c}) = h \in \{0, 1\}^{n_c}$ , then for any choice of  $\mu$ ,  $(Y_i)_{i \in I}$  is identically  $h$ .  
Similarly, if  $(\mu = 1 \text{ and } (\lambda^\ell > 0 \iff h_\ell = 1))$ , then  $(Y_i)_{i \in I}$  is identically  $h$  (with high probability as  $n_I \rightarrow \infty$ ).

## Theorem (Main result)

$I_2 := \{(i, j) \in I \times I \mid i \sim j, i < j\}$ ;  
 $\Omega \in \mathbb{R}^{2n_c \times 2n_c}$  is symmetric, positive definite;  $\alpha(\eta) = \eta^T \Omega \eta$  a quadratic form.

For triangular lattice:  $\Theta$  a compact subset of  $([0, 1]^{n_c} \setminus \{0, 1\}^{n_c}) \times [0, 1/5[$ .  
 (For the square lattice case, replace 1/5 with 1/3.)

When  $n_s$  is fixed and  $n_l \rightarrow \infty$ , then

$$\begin{aligned} \hat{\theta}_{n_s, n_l} &:= \arg \min_{\theta \in \Theta} \alpha \left( \begin{array}{c} \left( \frac{1}{n_l} \sum_{i \in I} \left( \mathcal{Y}_i^\ell - \frac{1}{n_s} \sum_{s=1}^{n_s} Y_i^{\ell, s} \right) \right)_{\ell \in \{1, \dots, n_c\}} \\ \left( \frac{1}{|I_2|} \sum_{(i, j) \in I_2} \left( \mathcal{Y}_i^\ell \mathcal{Y}_j^\ell - \frac{1}{n_s} \sum_{s=1}^{n_s} Y_i^{\ell, s} Y_j^{\ell, s} \right) \right)_{\ell \in \{1, \dots, n_c\}} \end{array} \right) \\ &= \arg \min_{\theta \in \Theta} \alpha \left( \begin{array}{c} \left( \bar{\mathcal{Y}}^\ell - \frac{1}{n_s} \sum_{s=1}^{n_s} \bar{Y}^{\ell, s} \right)_{\ell \in \{1, \dots, n_c\}} \\ \left( \bar{\mathcal{Z}}^\ell - \frac{1}{n_s} \sum_{s=1}^{n_s} \bar{Z}^{\ell, s} \right)_{\ell \in \{1, \dots, n_c\}} \end{array} \right) \end{aligned}$$

is strongly consistent.

# Goal: strong law of large numbers (SLLN)

Goal: as  $n_l \rightarrow \infty$ , for  $i \sim j$ , almost surely, uniformly in  $\theta \in \Theta$ ,

$$\frac{1}{n_l} \sum_{i \in l} Y_i^\ell \rightarrow \mathbb{E}_\theta Y_i^\ell$$

and

$$\frac{1}{|l_2|} \sum_{(i,j) \in l_2} Y_i^\ell Y_j^\ell \rightarrow \mathbb{E}_\theta [Y_i^\ell Y_j^\ell].$$

This would ensure that almost surely, uniformly in  $\theta \in \Theta$ ,

$$\alpha \left( \begin{array}{c} \left( \bar{Y}^\ell - \frac{1}{n_s} \sum_{S=1}^{n_s} \tilde{Y}^{\ell,S} \right)_{\ell \in \{1, \dots, n_c\}} \\ \left( \bar{Z}^\ell - \frac{1}{n_s} \sum_{S=1}^{n_s} \tilde{Z}^{\ell,S} \right)_{\ell \in \{1, \dots, n_c\}} \end{array} \right) \xrightarrow{n_l \rightarrow \infty} \alpha \left( \begin{array}{c} \left( \mathbb{E}_{\theta_0} Y_i^\ell - \mathbb{E}_\theta Y_i^\ell \right)_{\ell \in \{1, \dots, n_c\}} \\ \left( \mathbb{E}_{\theta_0} [Y_i^\ell Y_j^\ell] - \mathbb{E}_\theta [Y_i^\ell Y_j^\ell] \right)_{\ell \in \{1, \dots, n_c\}} \end{array} \right)$$

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# Strong law of large numbers (SLLN)



## Proposition (SLLN for our percolation model)

Let  $\Theta$  be a compact subset of  $[0, 1]^{n_c} \times [0, 1/5[$   
(triangular lattice:  $\mu < 1/5$ ; square lattice:  $\mu < 1/3$ ).  
If  $Y$  is generated with parameter value  $\theta \in \Theta$ , then

$$\frac{1}{n_l} \left( \sum_{i \in I} Y_i^\ell - \sum_{i \in I} \mathbb{E}_\theta Y_i^\ell \right) \xrightarrow[n_l \rightarrow \infty]{} 0$$

almost surely, uniformly in  $\theta \in \Theta$ .

## Proof

Let  $Y_i := Y_i^\ell$  for fixed  $\ell \in \{1, \dots, n_c\}$ .

Let  $a > 1$ . Define the lacunary sequence  $k_n := \lfloor a^n \rfloor$ .

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By Chebyshev's inequality for the  $\theta(k_n)$  for every  $k_n = [a^n]$  where the supremum on the compact set  $\Theta$  is achieved, for every  $\varepsilon > 0$ ,

$$\begin{aligned} \sum_{n=1}^{\infty} \mathbb{P} \left( \sup_{\theta \in \Theta} \left| \frac{S_{k_n} - \mathbb{E} S_{k_n}}{k_n} \right| > \varepsilon \right) &= \sum_{n=1}^{\infty} \mathbb{P} \left( \left| \frac{S_{k_n}(\theta(k_n)) - \mathbb{E}_{\theta(k_n)} S_{k_n}}{k_n} \right| > \varepsilon \right) \\ &\leq \sum_{n=1}^{\infty} \frac{\sup_{\theta \in \Theta} \text{Var} S_{k_n}}{\varepsilon^2 k_n^2} \\ &\leq \frac{1}{\varepsilon^2} \sum_{n=1}^{\infty} \frac{1}{k_n^2} \sup_{\theta \in \Theta} \sum_{i=1}^{k_n} \text{Var} Y_i \\ &\quad + \frac{1}{\varepsilon^2} \sum_{n=1}^{\infty} \frac{2}{k_n^2} \sup_{\theta \in \Theta} \sum_{1 \leq i < j \leq k_n} (\mathbb{E}[Y_i Y_j] - \mathbb{E} Y_i \mathbb{E} Y_j). \end{aligned}$$

If this is finite, then by the Borel–Cantelli lemma, as  $n \rightarrow \infty$ ,

$$\sup_{\theta \in \Theta} \left| \frac{S_{k_n} - \mathbb{E} S_{k_n}}{k_n} \right| \rightarrow 0 \quad \text{a.s.}$$

# Strong law of large numbers (SLLN)



$$\frac{1}{\varepsilon^2} \sum_{n=1}^{\infty} \frac{1}{k_n^2} \sup_{\theta \in \Theta} \sum_{i=1}^{k_n} \text{Var } Y_i + \frac{1}{\varepsilon^2} \sum_{n=1}^{\infty} \frac{2}{k_n^2} \sup_{\theta \in \Theta} \sum_{1 \leq i < j \leq k_n} (\mathbb{E}[Y_i Y_j] - \mathbb{E}Y_i \mathbb{E}Y_j)$$

## Lemma

If  $1 < a$ ,  $k_n = \lfloor a^n \rfloor$ , then

$$\sum_{n=1}^{\infty} \frac{1}{k_n} < \infty.$$

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As  $n \rightarrow \infty$ , it holds

$$\sup_{\theta \in \Theta} \left| \sum_{1 \leq i < j \leq n} (\mathbb{E}[Y_i Y_j] - \mathbb{E}Y_i \mathbb{E}Y_j) \right| = \mathcal{O}(n).$$

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# Quantifying (in)dependence of faraway vertices



Geoffrey Grimmett; *Percolation*, Springer, 1999.

**Probability space**  $(\{0, 1\}^S, \mathcal{F}, P)$  ( $|S| \leq \aleph_0$ );

**events**  $\mathcal{F}$ :  $\sigma$ -algebra generated by the finite-dimensional cylinder sets;

the **measure** is a product measure  $P = \prod_{s \in S} \nu_s$ ,

$\nu_s$  is given by  $(p(s))_{s \in S} \in [0, 1]^S$  via

$$\nu_s(\omega(s) = 1) = p(s), \quad \nu_s(\omega(s) = 0) = 1 - p(s)$$

for sample vectors  $(\omega(s))_{s \in S} \in \{0, 1\}^S$ .

A colour  $\ell \in \{1, \dots, n_c\}$  is already fixed.

Insert a loop edge for every vertex,  $p(s) := \lambda^\ell$ .

For edges of the lattice,  $p(s) := \mu$ .

An event  $A \in \mathcal{F}$  is **increasing** :  $\iff ((\omega \leq \omega', \omega \in A) \Rightarrow \omega' \in A)$ .

FKG inequality (Fortuin, Kasteleyn, Ginibre; 1971)

If  $A, B \in \mathcal{F}$  are increasing events, then  $P(A \cap B) \geq P(A)P(B)$ .

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$\nu_s$  is given by  $(p(s))_{s \in S} \in [0, 1]^S$  via

$$\nu_s(\omega(s) = 1) = p(s), \quad \nu_s(\omega(s) = 0) = 1 - p(s)$$

for sample vectors  $(\omega(s))_{s \in S} \in \{0, 1\}^S$ .

A colour  $\ell \in \{1, \dots, n_c\}$  is already fixed.

Insert a loop edge for every vertex,  $p(s) := \lambda^\ell$ .

For edges of the lattice,  $p(s) := \mu$ .

An event  $A \in \mathcal{F}$  is **increasing**:  $\iff ((\omega \leq \omega', \omega \in A) \Rightarrow \omega' \in A)$ .

## FKG inequality (Fortuin, Kasteleyn, Ginibre; 1971)

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# Some percolation theory

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Let  $e_1, e_2, \dots, e_N$  be  $N$  distinct edges,  $A, B \in \mathcal{F}$  two increasing events which depend on the states of these  $N$  edges  $\omega = (\omega(e_1), \dots, \omega(e_N))$  only.

$$J(\omega) := \{e_i \mid i \in \{1, \dots, N\}, \omega(e_i) = 1\}$$

For  $A, B$  increasing,  **$A$  and  $B$  occur disjointly**:

$A \circ B := \{\omega \in \{0, 1\}^S \mid \text{there exists an } H \subseteq J(\omega) \text{ such that } \omega' \text{ determined by } J(\omega') = H \text{ belongs to } A, \text{ and } \omega'' \text{ determined by } J(\omega'') = J(\omega) \setminus H \text{ belongs to } B\}$ .

$A \circ B$  is also increasing and  $A \circ B \subseteq A \cap B$ .

## BK inequality (van den Berg, Kesten; 1985)

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## Lemma

As  $n \rightarrow \infty$ , it holds

$$\sup_{\theta \in \Theta} \left| \sum_{1 \leq i < j \leq n} (E[Y_i Y_j] - EY_i EY_j) \right| = \mathcal{O}(n).$$

In the lattice graph extended with loop edges, the event  $\{Y_i = 1\}$  is *increasing*.

By the

**FKG inequality**,

If  $A, B \in \mathcal{F}$  are increasing events, then  $P(A \cap B) \geq P(A)P(B)$ .

$$E[Y_i Y_j] - EY_i EY_j = P(Y_i Y_j = 1) - P(Y_i = 1)P(Y_j = 1) \geq 0.$$

# Quantifying (in)dependence of faraway vertices

We have

$$E[Y_i Y_j] - E Y_i E Y_j = P(Y_i Y_j = 1) - P(Y_i = 1)P(Y_j = 1) \stackrel{\text{FKG}}{\geq} 0.$$

By the

**BK inequality**

If  $A, B \in \mathcal{F}$  are increasing events, then  $P(A \circ B) \leq P(A)P(B)$ .

$$\begin{aligned} P(Y_i Y_j = 1) - P(Y_i = 1)P(Y_j = 1) &= P(\{Y_i = 1\} \circ \{Y_j = 1\}) - P(Y_i = 1)P(Y_j = 1) \\ &\quad + P(\{Y_i Y_j = 1\} \setminus \{Y_i = 1\} \circ \{Y_j = 1\}) \\ &\stackrel{\text{BK}}{\leq} P(\{Y_i Y_j = 1\} \setminus \{Y_i = 1\} \circ \{Y_j = 1\}). \end{aligned}$$

Cooccurrence of  $\{Y_i = 1\}$  and  $\{Y_j = 1\}$  which is not disjoint is one where  $i$  and  $j$  are in the same component:

$$\{Y_i Y_j = 1\} \setminus \{Y_i = 1\} \circ \{Y_j = 1\} \subseteq \{i \leftrightarrow j\}.$$



# Quantifying (in)dependence of faraway vertices

Goal: on the triangular lattice, for every  $\varepsilon > 0$ , uniformly for  $\mu \in [0, 1/5 - \varepsilon]$ ,

$$\sum_{1 \leq i < j \leq n} P(i \leftrightarrow j) = \mathcal{O}(n)$$

$\mathcal{W}_k := \{\text{paths (i.e. self-avoiding walks) on the triangular lattice with length } k \text{ and beginning in a fixed vertex } i\}$

$$|\mathcal{W}_k| \leq 6 \times 5^{k-1}$$

$$E[\#\text{paths from } i] = \sum_{k=1}^{\infty} \sum_{\gamma \in \mathcal{W}_k} \mu^k = \sum_{k=1}^{\infty} |\mathcal{W}_k| \mu^k \leq \sum_{k=1}^{\infty} 6 \times 5^{k-1} \mu^k = 6\mu \frac{1}{1-5\mu} < \infty$$

By allowing any paths on the infinite lattice  $\supseteq I$ ,

$$\begin{aligned} \sum_{1 \leq i < j \leq n} P(i \leftrightarrow j) &\leq \sum_{1 \leq i \leq n} \sum_{k=1}^{\infty} \sum_{\substack{j \text{ is endpoint} \\ \text{of } \gamma \in \mathcal{W}_k}} \mu^k \\ &\leq \sum_{1 \leq i \leq n} \frac{6\mu}{1-5\mu} = \frac{6\mu}{1-5\mu} n. \end{aligned}$$

As  $n \rightarrow \infty$ , by the FKG and BK inequalities,

$$\sup_{\theta \in \Theta} \left| \sum_{1 \leq i < j \leq n} (E[Y_i Y_j] - E Y_i E Y_j) \right| = \mathcal{O}(n).$$

Hence for  $k_n = \lfloor a^n \rfloor$ ,

$$\sum_{n=1}^{\infty} P \left( \sup_{\theta \in \Theta} \left| \frac{S_{k_n} - ES_{k_n}}{k_n} \right| > \varepsilon \right) < \infty.$$

Further, consider  $k_n \leq n_l < k_{n+1}$ . Then

$$\frac{1}{n_l} \left( \sum_{i \in I} Y_i^\ell - \sum_{i \in I} E_\theta Y_i^\ell \right) \xrightarrow{n_l \rightarrow \infty} 0$$

almost surely, uniformly in  $\theta \in \Theta$ , where  $\Theta$  is a compact subset of  $[0, 1]^{n^c} \times [0, 1/5[$  ( $\mu < 1/3$  for the square lattice).

As  $n \rightarrow \infty$ , by the FKG and BK inequalities,

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almost surely, uniformly in  $\theta \in \Theta$ , where  $\Theta$  is a compact subset of  $[0, 1]^{n_c} \times [0, 1/5[$  ( $\mu < 1/3$  for the square lattice). **For the strong consistence of  $\hat{\theta}_{n_s, n_l}$ , repeat for**

$$\frac{1}{|I_2|} \left( \sum_{(i,j) \in I_2} Y_i^\ell Y_j^\ell - \sum_{(i,j) \in I_2} E_\theta [Y_i^\ell Y_j^\ell] \right) \xrightarrow{n_l \rightarrow \infty} 0.$$

$n_c = 3$  colours. Set  $\alpha(\eta) = \eta^T \Omega \eta$  by

$$\Omega = \text{diag}\left(\left(\bar{\mathcal{Y}}^1\right)^{-2}, \dots, \left(\bar{\mathcal{Y}}^{n_c}\right)^{-2}, \left(\bar{\mathcal{Z}}^1\right)^{-2}, \dots, \left(\bar{\mathcal{Z}}^{n_c}\right)^{-2}\right).$$

$$\alpha \left( \begin{array}{c} \left( \bar{\mathcal{Y}}^\ell - \frac{1}{n_s} \sum_{s=1}^{n_s} \bar{Y}^{\ell,s} \right)_{\ell \in \{1, \dots, n_c\}} \\ \left( \bar{\mathcal{Z}}^\ell - \frac{1}{n_s} \sum_{s=1}^{n_s} \bar{Z}^{\ell,s} \right)_{\ell \in \{1, \dots, n_c\}} \end{array} \right) \rightarrow \min$$

**Common random numbers** for different  $\theta = (\lambda^1, \dots, \lambda^{n_c}, \mu) \in \Theta$ .

Method 1:  $(U_i^{\ell,s}), (V_{ij}^s) \sim U[0, 1]$  ( $\ell \in \{1, \dots, n_c\}, s \in \{1, \dots, n_s\}, i \in I, (i, j) \in I_2$ )

$$X_i^{\ell,s} := \begin{cases} 1 & \text{if } U_i^{\ell,s} < \lambda^\ell, \\ 0 & \text{otherwise,} \end{cases} \quad \xi_{ij}^s := \begin{cases} 1 & \text{if } V_{ij}^s < \mu, \\ 0 & \text{otherwise.} \end{cases}$$

Method 2:  $(\sigma^{\ell,s}) \sim U[S_{n_\ell}], (\tau^s) \sim U[S_{|I_2|}]$  ( $\ell \in \{1, \dots, n_c\}, s \in \{1, \dots, n_s\}$ )

$$X_i^{\ell,s} := \begin{cases} 1 & \text{if } \sigma^{\ell,s}(i) \leq \lfloor \lambda^\ell n_\ell \rfloor, \\ 0 & \text{otherwise,} \end{cases} \quad \xi_{ij}^s := \begin{cases} 1 & \text{if } \tau^s((i, j)) \leq \lfloor \mu |I_2| \rfloor, \\ 0 & \text{otherwise.} \end{cases}$$

1  $n_I = 25 \times 25 = 625$  vertices ( $|I_2| = 1776$ )

$n_S$	$n_{\text{opt}}$	$\mu_{\text{max}}$	$\theta_0$	$\hat{\theta}_{n_S, n_I}^{(M1)}$	$d^{(M1)}$	$\hat{\theta}_{n_S, n_I}^{(M2)}$	$d^{(M2)}$	$\alpha_{\hat{\theta}_{n_S, n_I}^{(M2)}}$
10	10	0.1	0.1	0.1287	28.7%	0.1223	22.3%	0.0126
			0.05	0.0597	19.4%	0.0605	21%	
			0.07	0.0614	12.29%	0.0587	16.14%	
			0.06	0.0428	28.67%	0.0436	27.33%	

2  $n_I = 500 \times 500 = 250,000$  vertices ( $|I_2| = 748,001$ )

$n_S$	$n_{\text{opt}}$	$\mu_{\text{max}}$	$\theta_0$	$\hat{\theta}_{n_S, n_I}^{(M1)}$	$d^{(M1)}$	$\hat{\theta}_{n_S, n_I}^{(M2)}$	$d^{(M2)}$	$\alpha_{\hat{\theta}_{n_S, n_I}^{(M2)}}$
2	3	0.04	0.03	0.0295	1.67%	0.0293	2.33%	0.0011
			0.04	0.0402	0.5%	0.0401	0.25%	
			0.05	0.0522	4.4%	0.0520	4%	
			0.02	0.0192	4%	0.0195	2.5%	

$d$  is the relative bias:  $\left| 1 - \hat{\theta}_{n_S, n_I} / \theta_0 \right| \times 100\%$

- Model chosen: symmetric (undirected edges  $\xi_{ij}$ ); edges are independent Bernoulli variables; contamination propagates via a series of open edges.
- Method of simulated moments is strongly consistent as  $n_l \rightarrow \infty$  but  $n_s$  bounded.
- Unusual: sample is large but neither independent nor identically distributed.
- Proof method:
  - 1 The method of simulated moments (MSM) is strongly consistent.
  - 2 Proved a strong law of large numbers (SLLN) with weakly dependent variables.
  - 3 FKG and BK inequalities of percolation theory used to upper bound dependence (i.e. correlations bw. vertices).

- 1 Confidence intervals? Under regularity conditions (the estimator is continuously differentiable with respect to  $\theta$ ),  $\sqrt{n_I}(\hat{\theta}_{n_s, n_I} - \theta_0)$  is asymptotically normal with known limiting variance.
  - It is possible to choose  $\Omega$  optimally, i.e. to minimise this asymptotic variance.
- 2 Beyond estimating  $\mu$ , estimate the proportion of vertices which are in a non-trivial component.
- 3 Largest  $\mu$  for which SLLN holds? (cf.  $1/5$  for triangular,  $1/3$  for square lattice) Will this MSM work in the entire subcritical regime?
- 4 Maximum likelihood estimation; computing the probability of a configuration (and esp. of an *animal*).
- 5 Model fit? Locally positively correlated open edges might be needed; e.g. Ising model for the edges (increases degrees of freedom by 1).

- 1 Beck, Mélykúti; arXiv:1604.08908 [math.ST]
- 2 Gouriéroux, Monfort; *Simulation-based econometric methods*, Oxford University Press, Oxford, UK, 2002.
- 3 Gouriéroux, Monfort; Simulation based inference in models with heterogeneity. *Annales d'Économie et de Statistique*, 20–21:69–107, 1991.

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