

Latent Variable Models for the Analysis, Visualization and Prediction of Network and Nodal Attribute Data

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About Me

- With Jonty: Probabilistic methods for uncertainty assessment and quantification in natural hazards (floods, volcanoes, and earthquakes etc.).
- Models to cluster binary data with complex dependence structure
 - Gollini, I., and Murphy, T.B., (2013) "Mixture of Latent Trait Analyzers for Model-Based Clustering of Categorical Data", *Statistics and Computing*.
- Models for network data
- Use of Variational methods for fast approximate inference.

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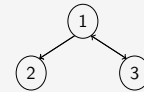
Outline

- Network Data
- Latent Space Models for Networks
 - Variational Inference
- Factor Analysis for Nodal Attributes
- Joint Model for Network and Nodal Attributes
- Latent Variable Models for Multiple Networks

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Notation

- N number of nodes of an observed network



- \mathbf{Y} ($N \times N$) adjacency matrix

$$y_{ij} = \begin{cases} 1 & \text{if there is an edge between node } i \text{ and } j \\ 0 & \text{otherwise} \end{cases}$$

- \mathbf{X} ($N \times M$) matrix of M nodal attributes.

- $\mathbf{z}_n \sim \mathcal{N}(\mathbf{0}, \sigma^2 \mathbf{1})$ D dimensional continuous latent variable

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Latent Space Model (LSM) for Networks

- Hoff *et al.* (2002) introduced a model that assumes that each node n has an unknown position \mathbf{z}_n in a D -dim *Euclidean latent space*.

$$p(\mathbf{Y}|\mathbf{Z}, \alpha) = \prod_{i \neq j} p(y_{ij} | \mathbf{z}_i, \mathbf{z}_j, \alpha) = \prod_{i \neq j} \frac{\exp(\alpha - |\mathbf{z}_i - \mathbf{z}_j|^2)^{y_{ij}}}{1 + \exp(\alpha - |\mathbf{z}_i - \mathbf{z}_j|^2)}$$

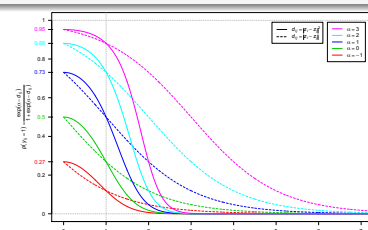
with $p(\alpha) = \mathcal{N}(\xi, \psi^2)$, $p(\mathbf{z}_n) \stackrel{iid}{=} \mathcal{N}(\mathbf{0}, \sigma^2 \mathbf{1}_D)$ and σ^2, ξ, ψ^2 are fixed parameters.

- The posterior distribution *cannot* be calculated analytically.
- NOTE: We propose to use is the *Squared Euclidean Distance*.

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Why Squared Euclidean Distance?

- It requires less approximation to be made in the estimation procedure.
- It allows to visualize more clearly the presence of potential clusters, giving a higher probability of a link between two close nodes in the latent space and lower probability to two nodes lying far away from each other.



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LSM for Networks – Variational Approach

- We fit the model using a *Variational inference approach* that is considerably quicker but less accurate than MCMC.

- The posterior probability of the unknown (\mathbf{Z}, α) is:

$$p(\mathbf{Z}, \alpha | \mathbf{Y}) = p(\mathbf{Y} | \mathbf{Z}, \alpha) p(\alpha) \prod_{n=1}^N p(\mathbf{z}_n) \times C$$

where C is the unknown normalising constant

- We propose a variational posterior $q(\mathbf{Z}, \alpha | \mathbf{Y})$ introducing variational parameters $\tilde{\xi}, \tilde{\psi}^2, \tilde{\mathbf{z}}_n, \tilde{\Sigma}$:

$$q(\mathbf{Z}, \alpha | \mathbf{Y}) = q(\alpha) \prod_{n=1}^N q(\mathbf{z}_n)$$

where $q(\alpha) = \mathcal{N}(\tilde{\xi}, \tilde{\psi}^2)$ and $q(\mathbf{z}_n) = \mathcal{N}(\tilde{\mathbf{z}}_n, \tilde{\Sigma})$.

Variational Approach

- The basic idea behind the variational approach is to find a lower bound of the log marginal likelihood $\log p(\mathbf{Y})$ by introducing the variational posterior distribution $q(\mathbf{Z}, \alpha | \mathbf{Y})$.

- This approach leads to minimize the Kulback-Leibler divergence between the variational posterior $q(\mathbf{Z}, \alpha | \mathbf{Y})$ and the true posterior $p(\mathbf{Z}, \alpha | \mathbf{Y})$:

$$\begin{aligned} \text{KL}[q(\mathbf{Z}, \alpha | \mathbf{Y}) || p(\mathbf{Z}, \alpha | \mathbf{Y})] &= - \int q(\mathbf{Z}, \alpha | \mathbf{Y}) \log \frac{p(\mathbf{Z}, \alpha | \mathbf{Y})}{q(\mathbf{Z}, \alpha | \mathbf{Y})} d(\mathbf{Z}, \alpha) \\ &= \int q(\mathbf{Z}, \alpha | \mathbf{Y}) \log \frac{p(\mathbf{Y}, \mathbf{Z}, \alpha)}{p(\mathbf{Y}) q(\mathbf{Z}, \alpha | \mathbf{Y})} d(\mathbf{Z}, \alpha) \\ &= \int q(\mathbf{Z}, \alpha | \mathbf{Y}) \log \frac{p(\mathbf{Y}, \mathbf{Z}, \alpha)}{q(\mathbf{Z}, \alpha | \mathbf{Y})} d(\mathbf{Z}, \alpha) - \log p(\mathbf{Y}) \end{aligned}$$

Variational Approach

- $\text{KL}[q(\mathbf{Z}, \alpha | \mathbf{Y}) || p(\mathbf{Z}, \alpha | \mathbf{Y})]$ divergence can be written as:

$$\begin{aligned} \text{KL}[q(\mathbf{Z}, \alpha | \mathbf{Y}) || p(\mathbf{Z}, \alpha | \mathbf{Y})] &= \text{KL}[q(\alpha) || p(\alpha)] + \sum_{i=1}^N \text{KL}[q(\mathbf{z}_i) || p(\mathbf{z}_i)] \\ &\quad - \mathbb{E}_{q(\mathbf{Z}, \alpha | \mathbf{Y})} [\log(p(\mathbf{Y} | \mathbf{Z}, \alpha))] \end{aligned}$$

$\mathbb{E}_{q(\mathbf{Z}, \alpha | \mathbf{Y})} [\log(p(\mathbf{Y} | \mathbf{Z}, \alpha))]$ is approximated using the Jensen's inequality:

$$\begin{aligned} \mathbb{E}_{q(\mathbf{Z}, \alpha | \mathbf{Y})} [\log(p(\mathbf{Y} | \mathbf{Z}, \alpha))] &= \sum_{i \neq j} y_{ij} \mathbb{E}_{q(\mathbf{Z}, \alpha | \mathbf{Y})} [\alpha - |\mathbf{z}_i - \mathbf{z}_j|^2] \\ &\quad - \mathbb{E}_{q(\mathbf{Z}, \alpha | \mathbf{Y})} [\log(1 + \exp(\alpha - |\mathbf{z}_i - \mathbf{z}_j|^2))] \\ &\leq \sum_{i \neq j} y_{ij} (\mathbb{E}_{q(\mathbf{Z}, \alpha | \mathbf{Y})} [\alpha - |\mathbf{z}_i - \mathbf{z}_j|^2]) \\ &\quad - \log(1 + \mathbb{E}_{q(\mathbf{Z}, \alpha | \mathbf{Y})} [\exp(\alpha - |\mathbf{z}_i - \mathbf{z}_j|^2)]) \end{aligned}$$

LSM for Networks – Variational Approach – EM Algorithm

- The EM algorithm at each $(i+1)$ th iteration:

E-Step Estimate $\tilde{\mathbf{z}}_n^{(i+1)}$ and $\tilde{\Sigma}^{(i+1)}$:

$$\mathcal{Q}(\Theta_{LSM}; \Theta_{LSM}^{(i)}) = \text{KL}[q(\mathbf{Z}, \alpha | \mathbf{Y}) || p(\mathbf{Z}, \alpha | \mathbf{Y})]$$

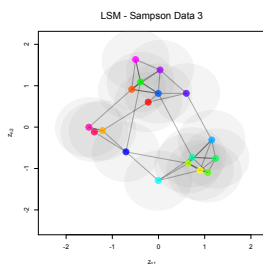
where $\Theta_{LSM} = (\tilde{\xi}, \tilde{\psi}^2)$.

M-Step Estimate $\tilde{\xi}$ and $\tilde{\psi}^2$:

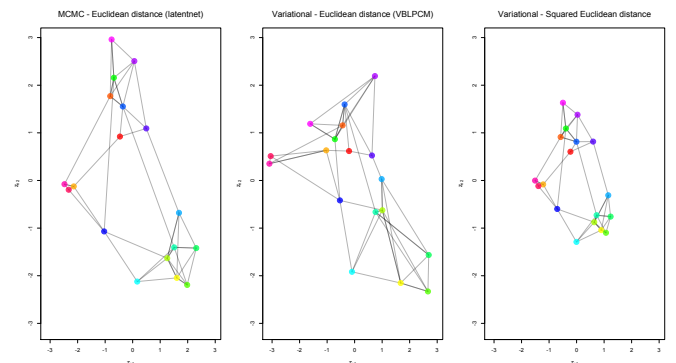
$$\Theta_{LSM}^{(i+1)} = \text{argmax} \mathcal{Q}(\Theta_{LSM}; \Theta_{LSM}^{(i)})$$

LSM – Monks Network

- Sampson (1969) recorded the social interactions among a group of $N = 18$ monks while being a resident in a New England monastery.
- The directed links of the network represent the liking relationships.



Comparison of Estimation Methods and Distance Metrics



Variational Inference VS MCMC

- Closed form posteriors
- Far faster than MCMC based methods
- In the absence of posterior dependence, the lower bound would match the log likelihood.
- As long as the posterior dependence is weak, the VA may be useful:
 - Larger networks
 - For starting point of MCMC algorithms
 - To explore the model space.
- Underestimates variances
- Difficult to assess how tight the lower bound is.
- Sensitive to starting values (local minima)

Network and Nodal Attributes

- The classical approach to incorporate nodal attributes in the LSM is:

$$p(\mathbf{Y}|\mathbf{Z}, \alpha, \mathbf{X}, \beta) = \prod_{i \neq j} \frac{\exp(\alpha + \beta^T \mathbf{x}_{ij} - |z_i - z_j|^2)^{y_{ij}}}{1 + \exp(\alpha + \beta^T \mathbf{x}_{ij} - |z_i - z_j|^2)}$$

- β and \mathbf{x}_{ij} are vectors of length M .
- This LSM contains only link covariate information \mathbf{x}_{ij} so it is not designed to deal with nodal attributes directly.
- This model assumes that the probability of a link depends on the nodal attributes (*social selection*)
- Sometimes the nodal attributes depend on the network links (*social influence*).
- We present a model where the network and the nodal attributes data mutually depend on each other.

Factor Analysis (FA) for Nodal Attributes

- Factor analysis (FA) (Spearman, 1904) is a useful technique to visualize continuous data, reducing the data dimensionality from M to D (where $D \ll M$) in order to explain the variability expressed by the correlation within the data.
- FA assumes that there is a continuous latent variable \mathbf{z}_n underlying the behavior of the continuous response variables given by an observation \mathbf{x}_n .

Factor Analysis (FA) for Nodal Attributes

$$\mathbf{z}_n \stackrel{iid}{\sim} \mathcal{N}(\mathbf{0}, \sigma^2 \mathbf{I}), \quad \text{and} \quad \boldsymbol{\varepsilon}_n \stackrel{iid}{\sim} \mathcal{N}(\mathbf{0}, \boldsymbol{\Psi}),$$

where $\boldsymbol{\Psi} = \text{diag}(\psi_1^2, \dots, \psi_M^2)$, and

$$\mathbf{x}_n = \boldsymbol{\mu} + \boldsymbol{\Lambda} \mathbf{z}_n + \boldsymbol{\varepsilon}_n$$

So,

$$p(\mathbf{x}_n | \mathbf{z}_n) \sim \mathcal{N}(\boldsymbol{\mu}, (\boldsymbol{\Lambda} \boldsymbol{\sigma})(\boldsymbol{\Lambda} \boldsymbol{\sigma})^T + \boldsymbol{\Psi})$$

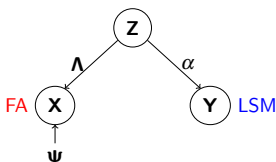
- The EM algorithm is used to find maximum likelihood estimate.

$$p(\mathbf{z}_n | \mathbf{x}_n) \sim \mathcal{N}(\hat{\mathbf{z}}_n, \hat{\boldsymbol{\Sigma}})$$

- Everything can be calculated analytically in closed form.

The joint model for Network and Nodal Attributes

- The probability of a node being connected with other nodes and the behaviour of nodal attributes are explained by the same latent variable.
- A continuous latent variable $\mathbf{z}_n \sim \mathcal{N}(0, \sigma^2 \mathbf{I}_D)$ summarizes the information given by both the network and the nodal attributes.
- Network \mathbf{Y} and nodal attributes \mathbf{x}_n are independent given the latent variable \mathbf{z}_n .



The joint model for Network and Nodal Attributes – Fit the model

- We assume that: $p(\mathbf{z}_n) \sim \mathcal{N}(\mathbf{0}, \sigma^2 \mathbf{I})$.
- The network data are modeled via LSM: $p(\mathbf{z}_n | \mathbf{Y}) \sim \mathcal{N}(\hat{\mathbf{z}}_n, \hat{\boldsymbol{\Sigma}})$.
- The nodal attributes are modeled via FA: $p(\mathbf{z}_n | \mathbf{x}_n) \sim \mathcal{N}(\hat{\mathbf{z}}_n, \hat{\boldsymbol{\Sigma}})$.

- Joint model:

$$p(\mathbf{z}_n | \mathbf{Y}, \mathbf{x}_n) \propto \frac{p(\mathbf{z}_n | \mathbf{Y}) p(\mathbf{z}_n | \mathbf{x}_n)}{p(\mathbf{z}_n)} \propto \mathcal{N}(\bar{\mathbf{z}}_n, \bar{\boldsymbol{\Sigma}})$$

where

$$\bar{\boldsymbol{\Sigma}} = \left[\hat{\boldsymbol{\Sigma}}^{-1} + \hat{\boldsymbol{\Sigma}}^{-1} - \frac{1}{\sigma^2} \mathbf{I}_D \right]^{-1} \quad \text{and} \quad \bar{\mathbf{z}}_n = \bar{\boldsymbol{\Sigma}} \left[\hat{\boldsymbol{\Sigma}}^{-1} \hat{\mathbf{z}}_n + \hat{\boldsymbol{\Sigma}}^{-1} \hat{\mathbf{z}}_n \right]$$

The joint model for Network and Nodal Attributes – EM Algorithm

E-Step Estimate $\bar{\Sigma}^{(i+1)}$ and $\bar{\mathbf{z}}_n^{(i+1)}$

$$\begin{aligned} \mathcal{L}(\Theta_{LSM}, \Theta_{FA}; \Theta_{LSM}^{(i)}, \Theta_{FA}^{(i)}) &= \\ &= \mathbb{E}_{p(\mathbf{Z}|\mathbf{Y}, \mathbf{X}; \Theta_{LSM}^{(i)}, \Theta_{FA}^{(i)})} [\log(p(\mathbf{Y}, \mathbf{Z}|\Theta_{LSM}))] + \\ &+ \mathbb{E}_{p(\mathbf{Z}|\mathbf{Y}, \mathbf{X}; \Theta_{LSM}^{(i)}, \Theta_{FA}^{(i)})} [\log(p(\mathbf{X}, \mathbf{Z}|\Theta_{FA}))] \end{aligned}$$

therefore,

$$\begin{aligned} \bar{\Sigma}^{(i+1)} &= \left[[\bar{\Sigma}^{(i+1)}]^{-1} + [\hat{\Sigma}^{(i+1)}]^{-1} - \frac{1}{\sigma^2} \mathbf{I}_D \right]^{-1} \\ \bar{\mathbf{z}}_n^{(i+1)} &= \bar{\Sigma}^{(i+1)} \left[[\bar{\Sigma}^{(i+1)}]^{-1} \bar{\mathbf{z}}_n^{(i+1)} + [\hat{\Sigma}^{(i+1)}]^{-1} \hat{\mathbf{z}}_n^{(i+1)} \right] \end{aligned}$$

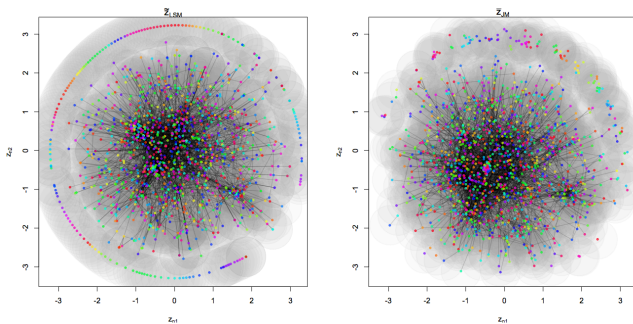
M-Step Update

$$(\Theta_{LSM}^{(i+1)}, \Theta_{FA}^{(i+1)}) = \operatorname{argmax} \mathcal{L}(\Theta_{LSM}, \Theta_{FA}; \Theta_{LSM}^{(i)}, \Theta_{FA}^{(i)})$$

The joint model for Network and Nodal Attributes – Yeast Proteins

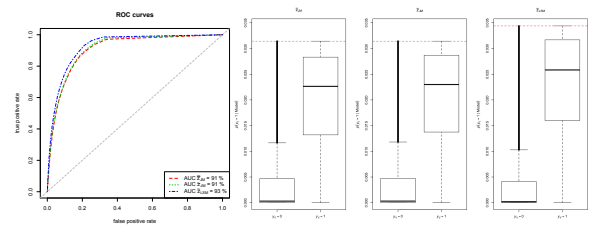
- The network is composed by $N=1500$ nodes representing the interaction between *Saccharomyces cerevisiae* (yeast) proteins.
- The nodal attributes consist of expression levels during yeast sporulation from $M=80$ experiments with *Saccharomyces cerevisiae* proteins.
- Factor analysis is an appropriate tool to visualize the M -dimensional expression data in a low dimensional latent space.
- The fixed parameters:
 - $\mathbf{z}_n \sim \mathcal{N}(\mathbf{0}, \mathbf{I}_2)$
 - $\xi = 0$
 - $\psi^2 = 2$

The joint model for Network and Nodal Attributes – Results



LSM positions (left) and LSJM positions (right).

The joint model for Network and Nodal Attributes – Performance



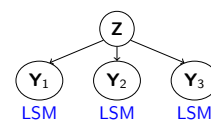
ROC curve (left) and Boxplot (right) of the estimated probabilities of a link for the true negatives and true positives.

Multiple Network Views

- In many applications the behaviour of the nodes is strongly shaped by the complex relation of many interactions.
- Longitudinal networks*: the links represent the same relation at different time points.
- Multiplex networks*: the links come from different kind of relations (eg genetic and physical etc.)

Joint Modelling of Multiple Network Views

- We have K networks on the same N nodes. We propose a model that merges the information given by all these networks.
- A continuous latent variable $\mathbf{z}_n \sim \mathcal{N}(\mathbf{0}, \sigma^2 \mathbf{I}_D)$ identifies the position of node n in a D -dimensional latent space.



Joint Modelling of Multiple Network Views – Model

- The probability of a link depends on the distance between two nodes in the latent space.

$$p(\mathbf{Y}_1, \dots, \mathbf{Y}_K | \mathbf{Z}, \alpha) = \prod_{k=1}^K \prod_{i \neq j} \frac{\exp(\alpha_k - |\mathbf{z}_i - \mathbf{z}_j|^2)^{y_{ijk}}}{1 + \exp(\alpha_k - |\mathbf{z}_i - \mathbf{z}_j|^2)}$$

- Variational Approach $k = 1, \dots, K$: $p(\mathbf{z}_n | \mathbf{Y}_k) \sim \mathcal{N}(\bar{\mathbf{z}}_{nk}, \bar{\Sigma}_k)$.
- Joining the two models:

$$p(\mathbf{z}_n | \mathbf{Y}_1, \dots, \mathbf{Y}_K; \Theta_1, \dots, \Theta_K) \propto \frac{\prod_{k=1}^K p(\mathbf{z}_n | \mathbf{Y}_k; \Theta_k)}{p(\mathbf{z}_n)^{K-1}} \\ \propto \mathcal{N}(\bar{\mathbf{z}}_n, \bar{\Sigma})$$

- where

$$\bar{\Sigma} = \left[\sum_{k=1}^K \bar{\Sigma}_k^{-1} - \frac{K-1}{\sigma^2} \mathbf{I}_D \right]^{-1} \text{ and } \bar{\mathbf{z}}_n = \bar{\Sigma} \left[\sum_{k=1}^K \bar{\Sigma}_k^{-1} \bar{\mathbf{z}}_{nk} \right]$$

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Joint Modelling of Multiple Network Views – EM Algorithm

E-Step Estimate the parameters of the joint posterior distribution $\bar{\Sigma}^{(i+1)}$ and $\bar{\mathbf{z}}_n^{(i+1)}$:

$$\mathcal{Q}(\Theta_1, \dots, \Theta_K; \Theta_1^{(i)}, \dots, \Theta_K^{(i)}) = \\ = \sum_{k=1}^K \mathbb{E}_{p(\mathbf{z}_n | \mathbf{Y}_1, \dots, \mathbf{Y}_K; \Theta_1^{(i)}, \dots, \Theta_K^{(i)})} [\log(p(\mathbf{Y}_k, \mathbf{Z} | \Theta_k))]$$

- We estimate the parameters $\bar{\mathbf{z}}_{nk}, \bar{\Sigma}_k$ of the posterior distribution $p(\mathbf{z}_n | \mathbf{Y}_k; \Theta_k)$ given each network k separately.
- We merge these estimates to find joint posterior distribution of the latent positions $\mathcal{N}(\bar{\mathbf{z}}_n, \bar{\Sigma})$.

M-Step Update the variational model parameters

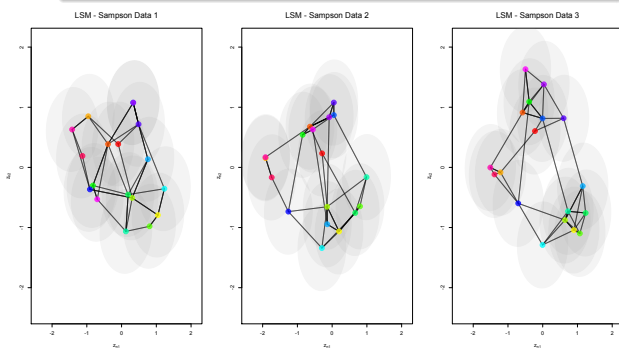
$$\bar{\Sigma}_1, \dots, \bar{\Sigma}_K; \bar{\Psi}_1^2, \dots, \bar{\Psi}_K^2:$$

$$(\Theta_1^{(i+1)}, \dots, \Theta_K^{(i+1)}) = \operatorname{argmax} \mathcal{Q}(\Theta_1, \dots, \Theta_K; \Theta_1^{(i)}, \dots, \Theta_K^{(i)}).$$

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LSJM – Monks Network (cont'd)

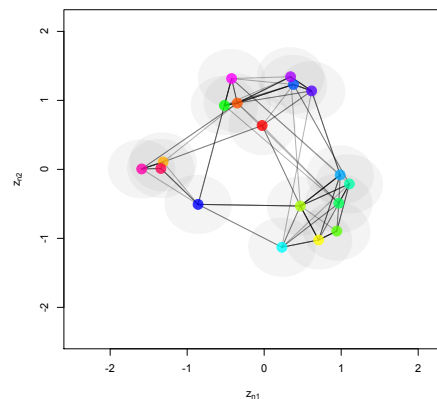
- We analyze the networks of liking relationship at $K = 3$ time points fitting the LSM to each network separately.



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LSJM – Monks Network – LSJM positions

LSJM - Sampson Data

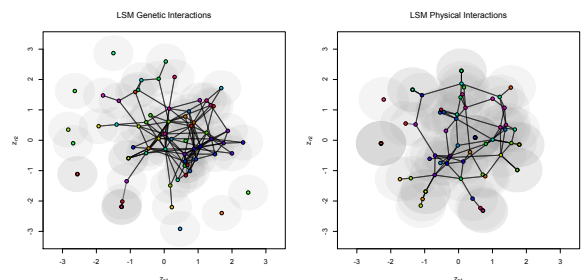


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LSJM – Protein-Protein Interactions

- $K = 2$ undirected networks formed by genetic and physical protein-protein interactions between $N = 67$ *Saccharomyces cerevisiae* proteins.
- The complex relational structure of this dataset has led to implementation of models aiming at describing the functional relationships between the observations.
- The data were downloaded from the Biological General Repository for Interaction Datasets (BioGRID) database.

LSJM – Protein-Protein Interactions – LSJM positions

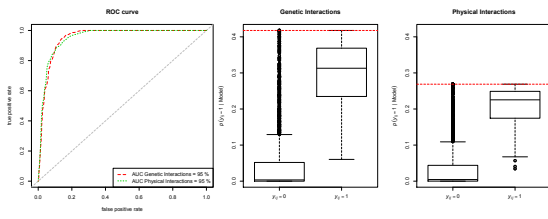


Latent posterior distributions fitting the LSM for the two networks separately.

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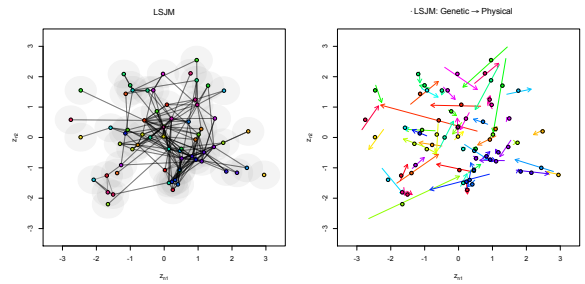
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LSJM – Protein-Protein Interactions – LSM ROC and BOX plots



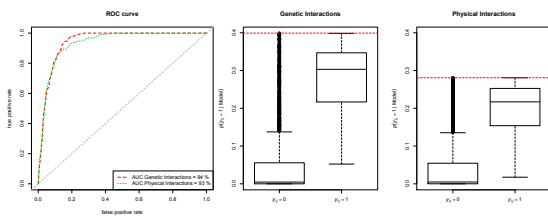
ROC curves and Boxplots of the estimated probabilities of a link for the true negatives and true positives.

LSJM – Protein-Protein Interactions – LSJM positions



- Left: $p(\mathbf{z}_n | \mathbf{Y}_1, \dots, \mathbf{Y}_K; \Theta_1, \dots, \Theta_K) \propto \mathcal{N}(\bar{\mathbf{z}}_n, \bar{\Sigma})$ fitting the LSJM
- Right: the dots represent the overall positions \bar{z}_n and the arrows connect the estimated position under each model $p(\mathbf{z}_n | \mathbf{Y}_k; \Theta_k) = \mathcal{N}(\tilde{z}_{nk}, \tilde{\Sigma}_k)$.

LSJM – Protein-Protein Interactions – LSJM ROC and BOX plots



ROC curves and Boxplots of the estimated probabilities of a link for the true negatives and true positives.

LSJM – Missing Links and Missing Nodes

- Missing (unobserved) links can be easily managed by the LSJM using the information given by all the network views.

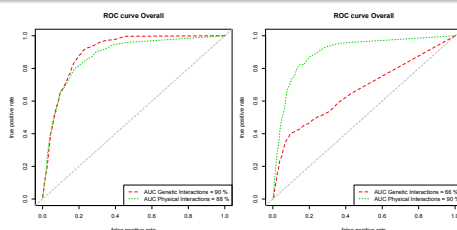
- To estimate the probability of the presence or absence of an edge we employ the posterior mean of the α_k and of the latent positions so that we get the following equation:

$$y_{ijk}^* = p(y_{ijk} = 1 | \bar{\mathbf{z}}_i, \bar{\mathbf{z}}_j, \tilde{\xi}_k) = \frac{\exp(\tilde{\xi}_k - |\bar{\mathbf{z}}_i - \bar{\mathbf{z}}_j|^2)}{1 + \exp(\tilde{\xi}_k - |\bar{\mathbf{z}}_i - \bar{\mathbf{z}}_j|^2)}$$

- If we want to infer whether to assign $y_{ijk} = 1$ or not, we need to introduce a threshold τ_k , and let $y_{ijk} = 1$ if $p(y_{ijk} = 1 | \bar{\mathbf{z}}_{ik}, \bar{\mathbf{z}}_{jk}, \tilde{\xi}_k) > \tau_k$.

LSJM – Protein-Protein Interactions – Missing Data

- To evaluate the link prediction we applied a 10-fold cross validation setting the 10% of the links to be missing at each time point.



ROC curves fitting a LSJM (left) and 2 single LSM (right)

LSJM – Protein-Protein Interactions – Missing Data

- Missing Links (10-fold cross validation):
 - LSJM: misclassification rate of 9% for the genetic interaction network, and 6% for the physical interaction network.
 - LSM: misclassification rate of 18% for the genetic interaction network, and 7% for the physical interaction network.

- Missing Nodes (10-fold cross validation):
 - LSJM: misclassification rate of 24% for the genetic interactions dataset and 20% for the physical interaction network.
 - LSM: *useless* since it would locate the nodes only relying on the prior information.

- Try to improve the predictions using a higher dimension for the latent variables.

Conclusions

The joint models are particularly useful to:

- Locate unconnected nodes/subgraphs in the latent space.
- Estimate missing links.
- Wide range of applications

Variational Bayes allows to deal with networks of thousands of nodes.

Possible extensions:

- Joint models for directed networks using the inner product instead of Euclidean distance.
- Joint models for categorical nodal attributes (LTA instead of FA).
- Joint models with clusters (LPCM, MFA, MLTA).
- Beyond binary networks: Rank and Count data.

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