

# Information-Theoretic Limits of Group Testing: Phase Transitions, Noisy Tests, and Partial Recovery

Jonathan Scarlett

*[jonathan.scarlett@epfl.ch](mailto:jonathan.scarlett@epfl.ch)*

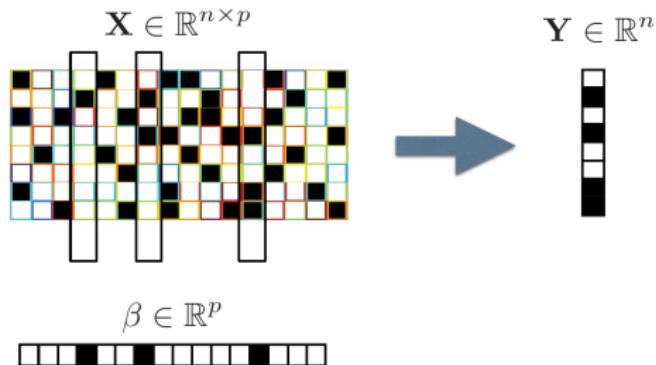
Laboratory for Information and Inference Systems (LIONS)  
École Polytechnique Fédérale de Lausanne (EPFL)  
Switzerland

Bristol Statistics Seminar  
[October, 2015]

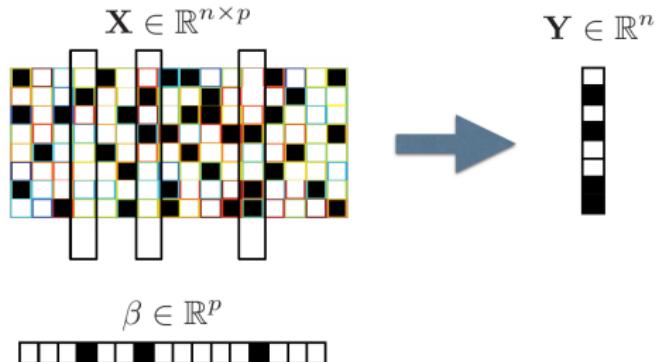
*Joint work with*  
Volkan Cevher @ LIONS



## Group testing

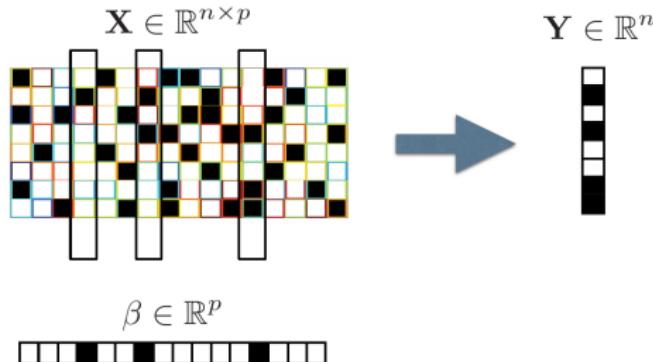


## Group testing



- ▶ In this talk:
  - ▶ Defective set  $S \sim \text{Uniform} \binom{[p]}{k}$
  - ▶ Measurement matrix  $\mathbf{X} \sim \prod_{i=1}^n \prod_{j=1}^p P_X(x_{i,j})$  with  $P_X \sim \text{Bernoulli}(\nu/k)$

## Group testing

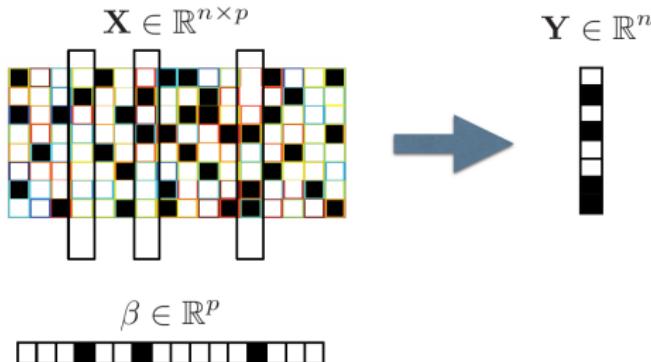


- ▶ In this talk:
  - ▶ Defective set  $S \sim \text{Uniform} \binom{p}{k}$
  - ▶ Measurement matrix  $\mathbf{X} \sim \prod_{i=1}^n \prod_{j=1}^p P_X(x_{i,j})$  with  $P_X \sim \text{Bernoulli}(\nu/k)$
- ▶ Perfect recovery
- ▶ Partial recovery

$$P_e := \mathbb{P}[\hat{S} \neq S]$$

$$P_e(d_{\max}) := \mathbb{P}\left[|S \setminus \hat{S}| > d_{\max} \cup |\hat{S} \setminus S| > d_{\max}\right]$$

## Group testing



- ▶ In this talk:
  - ▶ Defective set  $S \sim \text{Uniform}\binom{[p]}{k}$
  - ▶ Measurement matrix  $\mathbf{X} \sim \prod_{i=1}^n \prod_{j=1}^p P_X(x_{i,j})$  with  $P_X \sim \text{Bernoulli}(\nu/k)$
- ▶ Perfect recovery

$$P_e := \mathbb{P}[\hat{S} \neq S]$$

- ▶ Partial recovery

$$P_e(d_{\max}) := \mathbb{P}\left[|S \setminus \hat{S}| > d_{\max} \cup |\hat{S} \setminus S| > d_{\max}\right]$$

- ▶ **Goal:** Conditions on  $n$  for  $P_e \rightarrow 0$  or  $P_e(d_{\max}) \rightarrow 0$

# Noiseless Group Testing (Exact Recovery)

- ▶ Observation model

$$Y = \mathbf{1} \left\{ \bigcup_{i \in S} \{X_i = 1\} \right\}$$

- ▶ Bernoulli measurements, sparsity  $k = O(p^\theta)$

# Noiseless Group Testing (Exact Recovery)

- ▶ Observation model

$$Y = \mathbf{1} \left\{ \bigcup_{i \in S} \{X_i = 1\} \right\}$$

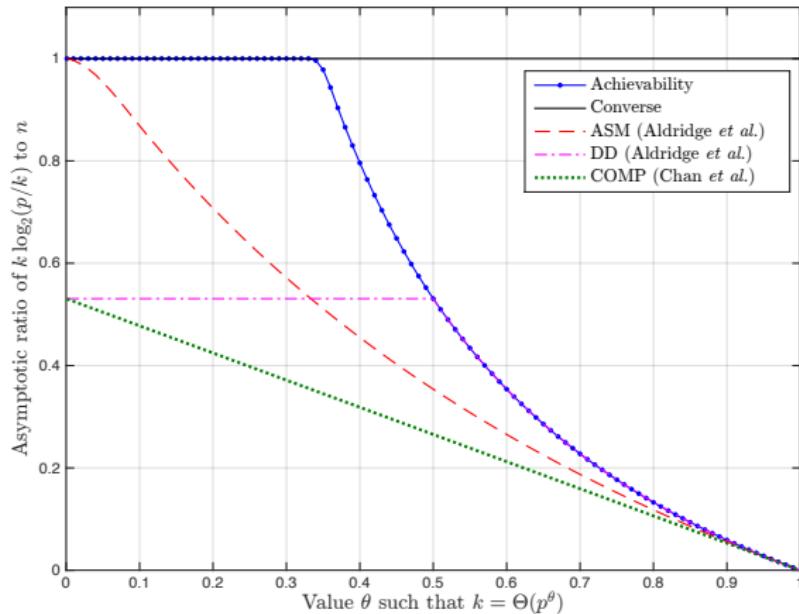
- ▶ Bernoulli measurements, sparsity  $k = O(p^\theta)$
- ▶ Sufficient for  $P_e \rightarrow 0$ :

$$n \geq \inf_{\nu > 0} \max \left\{ \frac{\theta}{e^{-\nu}\nu(1-\theta)}, \frac{1}{H_2(e^{-\nu})} \right\} \left( k \log \frac{p}{k} \right) (1 + \eta)$$

- ▶ Necessary for  $P_e \not\rightarrow 1$ :

$$n \geq \frac{k \log \frac{p}{k}}{\log 2} (1 - \eta)$$

## Noiseless Group Testing (Exact Recovery)



**Key Implication:** i.i.d. Bernoulli measurements are asymptotically as good as optimal adaptive measurements when  $k = O(p^{1/3})$ .

# Noisy Group Testing (Exact Recovery)

- ▶ Observation model

$$Y = \mathbf{1} \left\{ \bigcup_{i \in S} \{X_i = 1\} \right\} \oplus Z$$

where  $Z \sim \text{Bernoulli}(\rho)$

- ▶ Bernoulli measurements, sparsity  $k = O(p^\theta)$

# Noisy Group Testing (Exact Recovery)

- ▶ Observation model

$$Y = \mathbf{1} \left\{ \bigcup_{i \in S} \{X_i = 1\} \right\} \oplus Z$$

where  $Z \sim \text{Bernoulli}(\rho)$

- ▶ Bernoulli measurements, sparsity  $k = O(p^\theta)$

- ▶ Sufficient for  $P_e \rightarrow 0$ :

$$n \geq \inf_{\delta_2 \in (0,1)} \max \left\{ \zeta(\rho, \delta_2, \theta), \frac{1}{\log 2 - H_2(\rho)} \right\} \left( k \log \frac{p}{k} \right) (1 + \eta)$$

where

$$\zeta(\rho, \delta_2, \theta) := \frac{2}{\log 2} \max \left\{ \frac{2(1 + \frac{1}{3}\delta_2(1 - 2\rho)) \frac{\theta}{1-\theta}}{\delta_2^2(1 - 2\rho)^2}, \frac{\frac{1+2\theta}{1-\theta}}{(1 - 2\rho) \log \frac{1-\rho}{\rho}(1 - \delta_2)} \right\}.$$

# Noisy Group Testing (Exact Recovery)

- ▶ Observation model

$$Y = \mathbf{1} \left\{ \bigcup_{i \in S} \{X_i = 1\} \right\} \oplus Z$$

where  $Z \sim \text{Bernoulli}(\rho)$

- ▶ Bernoulli measurements, sparsity  $k = O(p^\theta)$

- ▶ Sufficient for  $P_e \rightarrow 0$ :

$$n \geq \inf_{\delta_2 \in (0,1)} \max \left\{ \zeta(\rho, \delta_2, \theta), \frac{1}{\log 2 - H_2(\rho)} \right\} \left( k \log \frac{p}{k} \right) (1 + \eta)$$

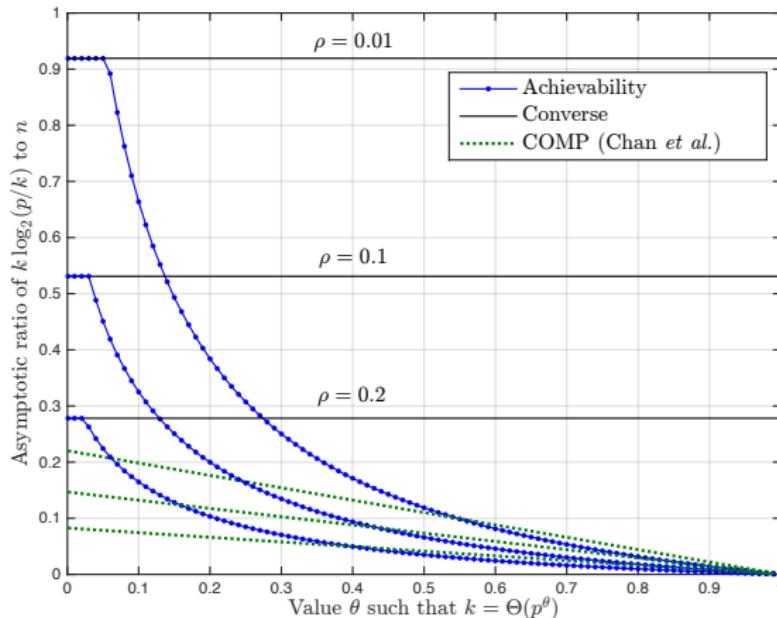
where

$$\zeta(\rho, \delta_2, \theta) := \frac{2}{\log 2} \max \left\{ \frac{2(1 + \frac{1}{3}\delta_2(1 - 2\rho)) \frac{\theta}{1-\theta}}{\delta_2^2(1 - 2\rho)^2}, \frac{\frac{1+2\theta}{1-\theta}}{(1 - 2\rho) \log \frac{1-\rho}{\rho}(1 - \delta_2)} \right\}.$$

- ▶ Necessary for  $P_e \not\rightarrow 1$ :

$$n \leq \frac{k \log \frac{p}{k}}{\log 2 - H_2(\rho)} (1 - \eta)$$

## Noisy Group Testing (Exact Recovery)



**Key Implication:** i.i.d. Bernoulli measurements are asymptotically as good as optimal adaptive measurements when  $k = O(p^\theta)$  for sufficiently small  $\theta$ .

## Partial Recovery

- ▶ Recovery criterion

$$P_e(d_{\max}) := \mathbb{P}\left[|S \setminus \hat{S}| > d_{\max} \cup |\hat{S} \setminus S| > d_{\max}\right]$$

where  $d_{\max} = \lfloor \alpha^* k \rfloor$  for some  $\alpha^* \in (0, 1)$

# Partial Recovery

- ▶ Recovery criterion

$$P_e(d_{\max}) := \mathbb{P}\left[|S \setminus \hat{S}| > d_{\max} \cup |\hat{S} \setminus S| > d_{\max}\right]$$

where  $d_{\max} = \lfloor \alpha^* k \rfloor$  for some  $\alpha^* \in (0, 1)$

- ▶ Sufficient for  $P_e(d_{\max}) \rightarrow 0$ :

$$n \geq \frac{k \log \frac{p}{k}}{\log 2 - H_2(\rho)} (1 + \eta)$$

- ▶ Necessary for  $P_e(d_{\max}) \not\rightarrow 1$ :

$$n \leq \frac{(1 - \alpha^*)k \log \frac{p}{k}}{\log 2 - H_2(\rho)} (1 - \eta)$$

# Partial Recovery

- ▶ Recovery criterion

$$P_e(d_{\max}) := \mathbb{P}\left[|S \setminus \hat{S}| > d_{\max} \cup |\hat{S} \setminus S| > d_{\max}\right]$$

where  $d_{\max} = \lfloor \alpha^* k \rfloor$  for some  $\alpha^* \in (0, 1)$

- ▶ Sufficient for  $P_e(d_{\max}) \rightarrow 0$ :

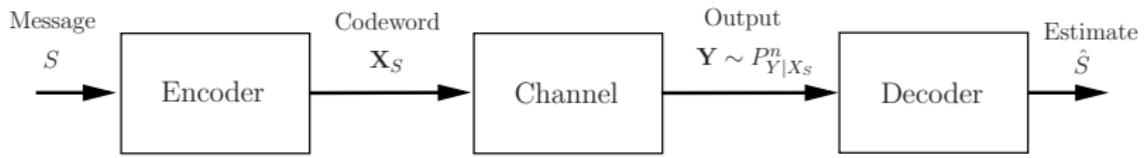
$$n \geq \frac{k \log \frac{p}{k}}{\log 2 - H_2(\rho)} (1 + \eta)$$

- ▶ Necessary for  $P_e(d_{\max}) \not\rightarrow 1$ :

$$n \leq \frac{(1 - \alpha^*)k \log \frac{p}{k}}{\log 2 - H_2(\rho)} (1 - \eta)$$

- ▶ **For small  $\theta$ , the reduction is at most a factor  $1 - \alpha^*$  asymptotically**

## Channel coding



e.g. see [Wainwright, 2009], [Atia and Saligrama, 2012], [Aksoylar et al., 2013]

# Thresholding Techniques for Channel Coding

- ▶ Mutual information

$$I(X; Y) := \sum_{x,y} P_{XY}(x, y) \log \frac{P_{Y|X}(y|x)}{P_Y(y)}$$

# Thresholding Techniques for Channel Coding

- ▶ Mutual information

$$I(X; Y) := \sum_{x,y} P_{XY}(x, y) \log \frac{P_{Y|X}(y|x)}{P_Y(y)}$$

- ▶ Information density

$$\imath(x; y) := \log \frac{P_{Y|X}(y|x)}{P_Y(y)}$$

# Thresholding Techniques for Channel Coding

- ▶ Mutual information

$$I(X; Y) := \sum_{x,y} P_{XY}(x,y) \log \frac{P_{Y|X}(y|x)}{P_Y(y)}$$

- ▶ Information density

$$\iota(x; y) := \log \frac{P_{Y|X}(y|x)}{P_Y(y)}$$

- ▶ Non-asymptotic channel coding bounds [Han, 2003]

$$P_e \leq \mathbb{P}\left[\iota^n(\mathbf{X}; \mathbf{Y}) \leq nR - \log \delta\right] + \delta$$

$$P_e \geq \mathbb{P}\left[\iota^n(\mathbf{X}; \mathbf{Y}) \leq nR + \log \delta\right] - \delta$$

# Non-Asymptotic Bounds

- ▶ Information density

$$\iota(x_{s_{\text{dif}}}; y|x_{s_{\text{eq}}}) := \log \frac{P_{Y|X_{s_{\text{dif}}} X_{s_{\text{eq}}}}(y|x_{s_{\text{dif}}}, x_{s_{\text{eq}}})}{P_{Y|X_{s_{\text{eq}}}}(y|x_{s_{\text{eq}}})}$$

where  $(s_{\text{dif}}, s_{\text{eq}})$  is a partition of  $s$

# Non-Asymptotic Bounds

- ▶ Information density

$$\iota(x_{s_{\text{dif}}}; y|x_{s_{\text{eq}}}) := \log \frac{P_{Y|X_{s_{\text{dif}}} X_{s_{\text{eq}}}}(y|x_{s_{\text{dif}}}, x_{s_{\text{eq}}})}{P_{Y|X_{s_{\text{eq}}}}(y|x_{s_{\text{eq}}})}$$

where  $(s_{\text{dif}}, s_{\text{eq}})$  is a partition of  $s$

- ▶ Achievability

$$P_e \leq \mathbb{P} \left[ \bigcup_{s_{\text{dif}}, s_{\text{eq}}} \left\{ \iota^n(\mathbf{X}_{s_{\text{dif}}}; \mathbf{Y}|\mathbf{X}_{s_{\text{eq}}}) \leq \log \binom{p-k}{|s_{\text{dif}}|} + \Delta \right\} \right] + \delta_1$$

# Non-Asymptotic Bounds

- ▶ Information density

$$\iota(x_{s_{\text{dif}}}; y|x_{s_{\text{eq}}}) := \log \frac{P_{Y|X_{s_{\text{dif}}} X_{s_{\text{eq}}}}(y|x_{s_{\text{dif}}}, x_{s_{\text{eq}}})}{P_{Y|X_{s_{\text{eq}}}}(y|x_{s_{\text{eq}}})}$$

where  $(s_{\text{dif}}, s_{\text{eq}})$  is a partition of  $s$

- ▶ Achievability

$$P_e \leq \mathbb{P} \left[ \bigcup_{s_{\text{dif}}, s_{\text{eq}}} \left\{ \iota^n(\mathbf{X}_{s_{\text{dif}}}; \mathbf{Y}|\mathbf{X}_{s_{\text{eq}}}) \leq \log \binom{p-k}{|s_{\text{dif}}|} + \Delta \right\} \right] + \delta_1$$

- ▶ Converse

$$P_e \geq \mathbb{P} \left[ \iota^n(\mathbf{X}_{s_{\text{dif}}}; \mathbf{Y}|\mathbf{X}_{s_{\text{eq}}}) \leq \log \binom{p-k+|s_{\text{dif}}|}{|s_{\text{dif}}|} + \log \delta_1 \right] - \delta_1$$

## Achievability Bound

- ▶ Decoder that searches for the unique set  $s \in \mathcal{S}$  such that

$$i^n(\mathbf{x}_{s_{\text{dif}}} ; \mathbf{y} | \mathbf{x}_{s_{\text{eq}}}) > \gamma_{|s_{\text{dif}}|}$$

for all  $(s_{\text{dif}}, s_{\text{eq}})$  of  $s$  with  $s_{\text{dif}} \neq \emptyset$ .

## Achievability Bound

- ▶ Decoder that searches for the unique set  $s \in \mathcal{S}$  such that

$$\vartheta^n(\mathbf{x}_{s_{\text{dif}}} ; \mathbf{y} | \mathbf{x}_{s_{\text{eq}}}) > \gamma_{|s_{\text{dif}}|}$$

for all  $(s_{\text{dif}}, s_{\text{eq}})$  of  $s$  with  $s_{\text{dif}} \neq \emptyset$ .

- ▶ Initial bound:

$$\begin{aligned} P_e \leq \mathbb{P} \left[ \bigcup_{(s_{\text{dif}}, s_{\text{eq}})} \left\{ \vartheta^n(\mathbf{X}_{s_{\text{dif}}} ; \mathbf{Y} | \mathbf{X}_{s_{\text{eq}}}) \leq \gamma_{|s_{\text{dif}}|} \right\} \right] \\ + \sum_{\bar{s} \in \mathcal{S} \setminus \{s\}} \mathbb{P} \left[ \vartheta^n(\mathbf{X}_{\bar{s} \setminus s} ; \mathbf{Y} | \mathbf{X}_{\bar{s} \cap s}) > \gamma_{|s_{\text{dif}}|} \right], \end{aligned}$$

## Achievability Bound

- ▶ Decoder that searches for the unique set  $s \in \mathcal{S}$  such that

$$\vartheta^n(\mathbf{x}_{s_{\text{dif}}}; \mathbf{y} | \mathbf{x}_{s_{\text{eq}}}) > \gamma_{|s_{\text{dif}}|}$$

for all  $(s_{\text{dif}}, s_{\text{eq}})$  of  $s$  with  $s_{\text{dif}} \neq \emptyset$ .

- ▶ Analysis of second term (with  $\ell := |\bar{s} \setminus s|$ ):

$$\begin{aligned} & \mathbb{P}\left[\vartheta^n(\mathbf{X}_{\bar{s} \setminus s}; \mathbf{Y} | \mathbf{X}_{\bar{s} \cap s}) > \gamma_\ell\right] \\ &= \sum_{\mathbf{x}_{\bar{s} \cap s}, \mathbf{x}_{\bar{s} \setminus s}, \mathbf{y}} P_X^{n \times (k-\ell)}(\mathbf{x}_{\bar{s} \cap s}) P_X^{n \times \ell}(\mathbf{x}_{\bar{s} \setminus s}) P_{Y|X_{s_{\text{eq}}}}^n(\mathbf{y} | \mathbf{x}_{\bar{s} \cap s}) \\ & \quad \times \mathbb{1} \left\{ \log \frac{P_{Y|X_{s_{\text{dif}}} X_{s_{\text{eq}}}}^n(\mathbf{y} | \mathbf{x}_{\bar{s} \setminus s}, \mathbf{x}_{\bar{s} \cap s})}{P_{Y|X_{s_{\text{eq}}}}^n(\mathbf{y} | \mathbf{x}_{\bar{s} \cap s})} > \gamma_\ell \right\} \end{aligned}$$

## Achievability Bound

- ▶ Decoder that searches for the unique set  $s \in \mathcal{S}$  such that

$$\iota^n(\mathbf{x}_{s_{\text{dif}}}; \mathbf{y} | \mathbf{x}_{s_{\text{eq}}}) > \gamma_{|s_{\text{dif}}|}$$

for all  $(s_{\text{dif}}, s_{\text{eq}})$  of  $s$  with  $s_{\text{dif}} \neq \emptyset$ .

- ▶ Analysis of second term (with  $\ell := |\bar{s} \setminus s|$ ):

$$\begin{aligned} & \mathbb{P}\left[\iota^n(\mathbf{X}_{\bar{s} \setminus s}; \mathbf{Y} | \mathbf{X}_{\bar{s} \cap s}) > \gamma_\ell\right] \\ &= \sum_{\mathbf{x}_{\bar{s} \cap s}, \mathbf{x}_{\bar{s} \setminus s}, \mathbf{y}} P_X^{n \times (k-\ell)}(\mathbf{x}_{\bar{s} \cap s}) P_X^{n \times \ell}(\mathbf{x}_{\bar{s} \setminus s}) P_{Y|X_{s_{\text{eq}}}}^n(\mathbf{y} | \mathbf{x}_{\bar{s} \cap s}) \\ & \quad \times \mathbb{1} \left\{ \log \frac{P_{Y|X_{s_{\text{dif}}} X_{s_{\text{eq}}}}^n(\mathbf{y} | \mathbf{x}_{\bar{s} \setminus s}, \mathbf{x}_{\bar{s} \cap s})}{P_{Y|X_{s_{\text{eq}}}}^n(\mathbf{y} | \mathbf{x}_{\bar{s} \cap s})} > \gamma_\ell \right\} \\ &\leq \sum_{\mathbf{x}_{\bar{s} \cap s}, \mathbf{x}_{\bar{s} \setminus s}, \mathbf{y}} P_X^{n \times (k-\ell)}(\mathbf{x}_{\bar{s} \cap s}) P_X^{n \times \ell}(\mathbf{x}_{\bar{s} \setminus s}) P_{Y|X_{s_{\text{dif}}} X_{s_{\text{eq}}}}^n(\mathbf{y} | \mathbf{x}_{\bar{s} \setminus s}, \mathbf{x}_{\bar{s} \cap s}) e^{-\gamma_\ell} \end{aligned}$$

## Achievability Bound

- ▶ Decoder that searches for the unique set  $s \in \mathcal{S}$  such that

$$\iota^n(\mathbf{x}_{s_{\text{dif}}}; \mathbf{y} | \mathbf{x}_{s_{\text{eq}}}) > \gamma_{|s_{\text{dif}}|}$$

for all  $(s_{\text{dif}}, s_{\text{eq}})$  of  $s$  with  $s_{\text{dif}} \neq \emptyset$ .

- ▶ Analysis of second term (with  $\ell := |\bar{s} \setminus s|$ ):

$$\begin{aligned} & \mathbb{P}\left[\iota^n(\mathbf{X}_{\bar{s} \setminus s}; \mathbf{Y} | \mathbf{X}_{\bar{s} \cap s}) > \gamma_\ell\right] \\ &= \sum_{\mathbf{x}_{\bar{s} \cap s}, \mathbf{x}_{\bar{s} \setminus s}, \mathbf{y}} P_X^{n \times (k-\ell)}(\mathbf{x}_{\bar{s} \cap s}) P_X^{n \times \ell}(\mathbf{x}_{\bar{s} \setminus s}) P_{Y|X_{s_{\text{eq}}}}^n(\mathbf{y} | \mathbf{x}_{\bar{s} \cap s}) \\ & \quad \times \mathbb{1} \left\{ \log \frac{P_{Y|X_{s_{\text{dif}}} X_{s_{\text{eq}}}}^n(\mathbf{y} | \mathbf{x}_{\bar{s} \setminus s}, \mathbf{x}_{\bar{s} \cap s})}{P_{Y|X_{s_{\text{eq}}}}^n(\mathbf{y} | \mathbf{x}_{\bar{s} \cap s})} > \gamma_\ell \right\} \\ &\leq \sum_{\mathbf{x}_{\bar{s} \cap s}, \mathbf{x}_{\bar{s} \setminus s}, \mathbf{y}} P_X^{n \times (k-\ell)}(\mathbf{x}_{\bar{s} \cap s}) P_X^{n \times \ell}(\mathbf{x}_{\bar{s} \setminus s}) P_{Y|X_{s_{\text{dif}}} X_{s_{\text{eq}}}}^n(\mathbf{y} | \mathbf{x}_{\bar{s} \setminus s}, \mathbf{x}_{\bar{s} \cap s}) e^{-\gamma_\ell} \\ &= e^{-\gamma_\ell} \end{aligned}$$

## Achievability Bound

- ▶ Decoder that searches for the unique set  $s \in \mathcal{S}$  such that

$$\vartheta^n(\mathbf{x}_{s_{\text{dif}}}; \mathbf{y} | \mathbf{x}_{s_{\text{eq}}}) > \gamma_{|s_{\text{dif}}|}$$

for all  $(s_{\text{dif}}, s_{\text{eq}})$  of  $s$  with  $s_{\text{dif}} \neq \emptyset$ .

- ▶ After some re-arrangements and choosing  $\{\gamma_\ell\}$  appropriately,

$$P_e \leq \mathbb{P} \left[ \bigcup_{s_{\text{dif}}, s_{\text{eq}}} \left\{ \vartheta^n(\mathbf{X}_{s_{\text{dif}}}; \mathbf{Y} | \mathbf{X}_{s_{\text{eq}}}) \leq \log \binom{p-k}{|s_{\text{dif}}|} + \Delta \right\} \right] + \delta_1$$

## Converse Bound

- ▶ Consider a genie that reveals part of the defective set,  $S_{\text{eq}} \subseteq S$ , to the decoder.  
The decoder is left to estimate  $S_{\text{dif}} := S \setminus S_{\text{eq}}$ .
- ▶ Starting point:

$$P_e(s_{\text{eq}}) \geq \mathbb{P}[\mathcal{A}(s_{\text{eq}})] - \mathbb{P}[\mathcal{A}(s_{\text{eq}}) \cap \text{no error}],$$

where

$$\mathcal{A}(s_{\text{eq}}) = \left\{ \vartheta^n(\mathbf{X}_{S_{\text{dif}}}; \mathbf{Y} | \mathbf{X}_{s_{\text{eq}}}) \leq \gamma \right\}.$$

## Converse Bound

- ▶ Consider a genie that reveals part of the defective set,  $S_{\text{eq}} \subseteq S$ , to the decoder. The decoder is left to estimate  $S_{\text{dif}} := S \setminus S_{\text{eq}}$ .
- ▶ Analysis of second term:

$$\mathbb{P}[\mathcal{A}(s_{\text{eq}}) \cap \text{no error}]$$

$$\begin{aligned} &= \sum_{s_{\text{dif}} \in \mathcal{S}_{\text{dif}}(s_{\text{eq}})} \frac{1}{\binom{p-k+\ell}{\ell}} \sum_{(\mathbf{x}, \mathbf{y}) \in \mathcal{D}(s_{\text{dif}} | s_{\text{eq}})} P_X^{n \times p}(\mathbf{x}) \\ &\quad \times P_{Y|X_{s_{\text{dif}}} X_{s_{\text{eq}}}}^n(\mathbf{y} | \mathbf{x}_{s_{\text{dif}}}, \mathbf{x}_{s_{\text{eq}}}) \mathbb{1} \left\{ \log \frac{P_{Y|X_{s_{\text{dif}}} X_{s_{\text{eq}}}}^n(\mathbf{y} | \mathbf{x}_{s_{\text{dif}}}, \mathbf{x}_{s_{\text{eq}}})}{P_{Y|X_{s_{\text{eq}}}}^n(\mathbf{y} | \mathbf{x}_{s_{\text{eq}}})} \leq \gamma \right\} \end{aligned}$$

## Converse Bound

- ▶ Consider a genie that reveals part of the defective set,  $S_{\text{eq}} \subseteq S$ , to the decoder. The decoder is left to estimate  $S_{\text{dif}} := S \setminus S_{\text{eq}}$ .
- ▶ Analysis of second term:

$$\mathbb{P}[\mathcal{A}(s_{\text{eq}}) \cap \text{no error}]$$

$$\begin{aligned} &= \sum_{s_{\text{dif}} \in \mathcal{S}_{\text{dif}}(s_{\text{eq}})} \frac{1}{\binom{p-k+\ell}{\ell}} \sum_{(\mathbf{x}, \mathbf{y}) \in \mathcal{D}(s_{\text{dif}} | s_{\text{eq}})} P_X^{n \times p}(\mathbf{x}) \\ &\quad \times P_{Y|X_{s_{\text{dif}}} X_{s_{\text{eq}}}}^n(\mathbf{y} | \mathbf{x}_{s_{\text{dif}}}, \mathbf{x}_{s_{\text{eq}}}) \mathbb{1} \left\{ \log \frac{P_{Y|X_{s_{\text{dif}}} X_{s_{\text{eq}}}}^n(\mathbf{y} | \mathbf{x}_{s_{\text{dif}}}, \mathbf{x}_{s_{\text{eq}}})}{P_{Y|X_{s_{\text{eq}}}}^n(\mathbf{y} | \mathbf{x}_{s_{\text{eq}}})} \leq \gamma \right\} \\ &\leq \frac{1}{\binom{p-k+\ell}{\ell}} \sum_{s_{\text{dif}} \in \mathcal{S}_{\text{dif}}(s_{\text{eq}})} \sum_{(\mathbf{x}, \mathbf{y}) \in \mathcal{D}(s_{\text{dif}} | s_{\text{eq}})} P_X^{n \times p}(\mathbf{x}) P_{Y|X_{s_{\text{eq}}}}^n(\mathbf{y} | \mathbf{x}_{s_{\text{eq}}}) e^\gamma \end{aligned}$$

## Converse Bound

- ▶ Consider a genie that reveals part of the defective set,  $S_{\text{eq}} \subseteq S$ , to the decoder. The decoder is left to estimate  $S_{\text{dif}} := S \setminus S_{\text{eq}}$ .
- ▶ Analysis of second term:

$$\mathbb{P}[\mathcal{A}(s_{\text{eq}}) \cap \text{no error}]$$

$$\begin{aligned} &= \sum_{s_{\text{dif}} \in \mathcal{S}_{\text{dif}}(s_{\text{eq}})} \frac{1}{\binom{p-k+\ell}{\ell}} \sum_{(\mathbf{x}, \mathbf{y}) \in \mathcal{D}(s_{\text{dif}} | s_{\text{eq}})} P_X^{n \times p}(\mathbf{x}) \\ &\quad \times P_{Y|X_{s_{\text{dif}}} X_{s_{\text{eq}}}}^n(\mathbf{y} | \mathbf{x}_{s_{\text{dif}}}, \mathbf{x}_{s_{\text{eq}}}) \mathbb{1} \left\{ \log \frac{P_{Y|X_{s_{\text{dif}}} X_{s_{\text{eq}}}}^n(\mathbf{y} | \mathbf{x}_{s_{\text{dif}}}, \mathbf{x}_{s_{\text{eq}}})}{P_{Y|X_{s_{\text{eq}}}}^n(\mathbf{y} | \mathbf{x}_{s_{\text{eq}}})} \leq \gamma \right\} \\ &\leq \frac{1}{\binom{p-k+\ell}{\ell}} \sum_{s_{\text{dif}} \in \mathcal{S}_{\text{dif}}(s_{\text{eq}})} \sum_{(\mathbf{x}, \mathbf{y}) \in \mathcal{D}(s_{\text{dif}} | s_{\text{eq}})} P_X^{n \times p}(\mathbf{x}) P_{Y|X_{s_{\text{eq}}}}^n(\mathbf{y} | \mathbf{x}_{s_{\text{eq}}}) e^\gamma \\ &= \frac{e^\gamma}{\binom{p-k+\ell}{\ell}}. \end{aligned}$$

## Converse Bound

- ▶ Consider a genie that reveals part of the defective set,  $S_{\text{eq}} \subseteq S$ , to the decoder. The decoder is left to estimate  $S_{\text{dif}} := S \setminus S_{\text{eq}}$ .
- ▶ After some re-arrangements and choosing  $\{\gamma_\ell\}$  appropriately,

$$P_e \geq \mathbb{P} \left[ I^n(\mathbf{X}_{S_{\text{dif}}}; \mathbf{Y} | \mathbf{X}_{S_{\text{eq}}}) \leq \log \left( \frac{p - k + |S_{\text{dif}}|}{|S_{\text{dif}}|} \right) + \log \delta_1 \right] - \delta_1$$

# Non-Asymptotic Bounds

- ▶ Information density

$$\iota(x_{s_{\text{dif}}}; y|x_{s_{\text{eq}}}) := \log \frac{P_{Y|X_{s_{\text{dif}}} X_{s_{\text{eq}}}}(y|x_{s_{\text{dif}}}, x_{s_{\text{eq}}})}{P_{Y|X_{s_{\text{eq}}}}(y|x_{s_{\text{eq}}})}$$

where  $(s_{\text{dif}}, s_{\text{eq}})$  is a partition of  $s$

- ▶ Achievability

$$P_e \leq \mathbb{P} \left[ \bigcup_{s_{\text{dif}}, s_{\text{eq}}} \left\{ \iota^n(\mathbf{X}_{s_{\text{dif}}}; \mathbf{Y}|\mathbf{X}_{s_{\text{eq}}}) \leq \log \binom{p-k}{|s_{\text{dif}}|} + \Delta \right\} \right] + \delta_1$$

- ▶ Converse

$$P_e \geq \mathbb{P} \left[ \iota^n(\mathbf{X}_{s_{\text{dif}}}; \mathbf{Y}|\mathbf{X}_{s_{\text{eq}}}) \leq \log \binom{p-k+|s_{\text{dif}}|}{|s_{\text{dif}}|} + \log \delta_1 \right] - \delta_1$$

# Application Techniques

- ▶ General steps for application

1. Bound the tail probabilities  $\mathbb{P}\left[\iota^n(\mathbf{X}_{s_{\text{dif}}}; \mathbf{Y} | \mathbf{X}_{s_{\text{eq}}}) \leq \mathbb{E}[\iota^n] \pm n\delta\right]$

# Application Techniques

- ▶ General steps for application

1. Bound the tail probabilities  $\mathbb{P}\left[\iota^n(\mathbf{X}_{s_{\text{dif}}}; \mathbf{Y} | \mathbf{X}_{s_{\text{eq}}}) \leq \mathbb{E}[\iota^n] \pm n\delta\right]$
2. Control and simplify the remainder terms

# Application Techniques

- ▶ General steps for application
  1. Bound the tail probabilities  $\mathbb{P}\left[\iota^n(\mathbf{X}_{s_{\text{dif}}}; \mathbf{Y} | \mathbf{X}_{s_{\text{eq}}}) \leq \mathbb{E}[\iota^n] \pm n\delta\right]$
  2. Control and simplify the remainder terms
- ▶ General form of corollaries:  $P_e \rightarrow 0$  if

$$n \geq \max_{(s_{\text{dif}}, s_{\text{eq}})} \frac{\log \binom{p-k}{|s_{\text{dif}}|}}{I(X_{s_{\text{dif}}}; Y | X_{s_{\text{eq}}})} (1 + \eta)$$

and  $P_e \rightarrow 1$  if

$$n \leq \max_{(s_{\text{dif}}, s_{\text{eq}})} \frac{\log \binom{p-k+|s_{\text{dif}}|}{|s_{\text{dif}}|}}{I(X_{s_{\text{dif}}}; Y | X_{s_{\text{eq}}})} (1 - \eta)$$

## Concentration Inequalities

- ▶ Noiseless and noisy cases (*This one alone is enough for partial recovery!*):

$$\mathbb{P}\left[\left|\vartheta^n(\mathbf{X}_{\text{sdif}}; \mathbf{Y} | \mathbf{X}_{\text{seq}}) - nI(\ell)\right| \geq n\delta\right] \leq 2 \exp\left(-\frac{\delta^2 n}{4(8 + \delta)}\right)$$

- ▶ Proved using Bernstein's inequality
- ▶ The only property used is that  $|\mathcal{Y}| = 2$

# Concentration Inequalities

- ▶ Noiseless and noisy cases (*This one alone is enough for partial recovery!*):

$$\mathbb{P}\left[\left|\imath^n(\mathbf{X}_{s_{\text{dif}}}; \mathbf{Y} | \mathbf{X}_{s_{\text{eq}}}) - nI(\ell)\right| \geq n\delta\right] \leq 2 \exp\left(-\frac{\delta^2 n}{4(8 + \delta)}\right)$$

- ▶ Proved using Bernstein's inequality
- ▶ The only property used is that  $|\mathcal{Y}| = 2$

- ▶ Noiseless case:

$$\mathbb{P}\left[\imath^n \leq nI(\ell)(1 - \delta_2)\right] \leq \exp\left(-n\frac{\ell}{k}e^{-\nu}\nu\left((1 - \delta_2)\log(1 - \delta_2) + \delta_2\right)(1 - \epsilon)\right)$$

- ▶ Proved by writing  $\imath^n(\mathbf{X}_{s_{\text{dif}}}; \mathbf{Y} | \mathbf{X}_{s_{\text{eq}}})$  in terms of Binomial random variables, then applying Binomial tail bounds

# Concentration Inequalities

- ▶ Noiseless and noisy cases (*This one alone is enough for partial recovery!*):

$$\mathbb{P}\left[\left|\imath^n(\mathbf{X}_{s_{\text{dif}}}; \mathbf{Y} | \mathbf{X}_{s_{\text{eq}}}) - nI(\ell)\right| \geq n\delta\right] \leq 2 \exp\left(-\frac{\delta^2 n}{4(8 + \delta)}\right)$$

- ▶ Proved using Bernstein's inequality
- ▶ The only property used is that  $|\mathcal{Y}| = 2$

- ▶ Noiseless case:

$$\mathbb{P}\left[\imath^n \leq nI(\ell)(1 - \delta_2)\right] \leq \exp\left(-n\frac{\ell}{k}e^{-\nu}\nu\left((1 - \delta_2)\log(1 - \delta_2) + \delta_2\right)(1 - \epsilon)\right)$$

- ▶ Proved by writing  $\imath^n(\mathbf{X}_{s_{\text{dif}}}; \mathbf{Y} | \mathbf{X}_{s_{\text{eq}}})$  in terms of Binomial random variables, then applying Binomial tail bounds

- ▶ Noisy case with crossover probability  $\rho$ :

$$\mathbb{P}\left[\imath^n \leq nI(\ell)(1 - \delta_2)\right] \leq \exp\left(-n\frac{\ell}{k}e^{-\nu}\nu\left(\frac{\delta_2^2(1 - 2\rho)^2}{2(1 + \frac{1}{3}\delta_2(1 - 2\rho))}\right)(1 - \epsilon)\right)$$

- ▶ Proved using Bennet's inequality – may be somewhat crude

# Recap of Results

- ▶ **Noiseless case (exact recovery):**

- ▶ Sufficient for  $P_e \rightarrow 0$ :

$$n \geq \inf_{\nu > 0} \max \left\{ \frac{\theta}{e^{-\nu} \nu (1 - \theta)}, \frac{1}{H_2(e^{-\nu})} \right\} \left( k \log \frac{p}{k} \right) (1 + \eta)$$

- ▶ Necessary for  $P_e \not\rightarrow 1$ :

$$n \geq \frac{k \log \frac{p}{k}}{\log 2} (1 - \eta)$$

- ▶ **Noisy case (exact recovery):**

- ▶ Sufficient for  $P_e \rightarrow 0$ :

$$n \geq \inf_{\delta_2 \in (0, 1)} \max \left\{ \zeta(\rho, \delta_2, \theta), \frac{1}{\log 2 - H_2(\rho)} \right\} \left( k \log \frac{p}{k} \right) (1 + \eta).$$

- ▶ Necessary for  $P_e \not\rightarrow 1$ :

$$n \leq \frac{k \log \frac{p}{k}}{\log 2 - H_2(\rho)} (1 - \eta)$$

- ▶ **Noiseless and noisy cases (partial recovery):**

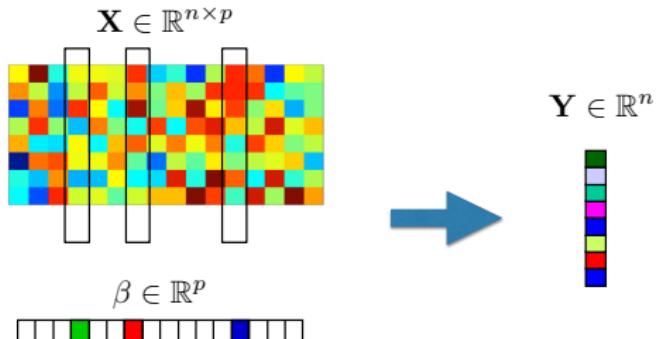
- ▶ Sufficient for  $P_e(d_{\max}) \rightarrow 0$ :

$$n \geq \frac{k \log \frac{p}{k}}{\log 2 - H_2(\rho)} (1 + \eta)$$

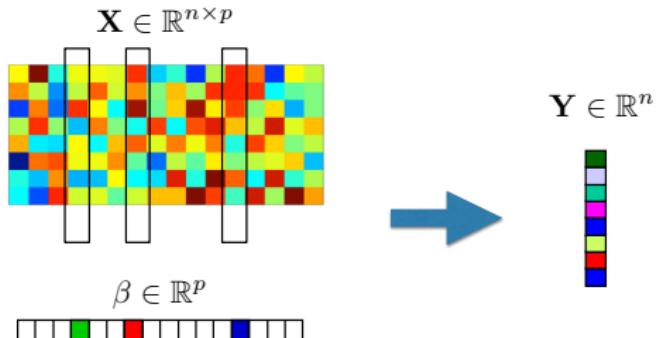
- ▶ Necessary for  $P_e(d_{\max}) \not\rightarrow 1$ :

$$n \leq \frac{(1 - \alpha^*) k \log \frac{p}{k}}{\log 2 - H_2(\rho)} (1 - \eta)$$

## More General Support Recovery

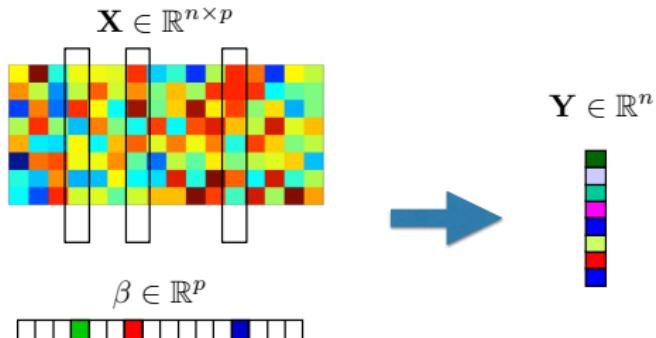


# More General Support Recovery



- ▶ General probabilistic models
  - ▶ Support  $S \sim \text{Uniform} \binom{p}{k}$
  - ▶ Measurement matrix  $\mathbf{X} \sim \prod_{i=1}^n \prod_{j=1}^p P_X(x_{i,j})$
  - ▶ Non-zero entries  $\beta_S \sim P_{\beta_S}$
  - ▶ Observations  $(\mathbf{Y}|\mathbf{X}, \beta) \sim P_{\mathbf{Y}|\mathbf{X}_S \beta_S}$

# More General Support Recovery



- ▶ General probabilistic models

- ▶ Support  $S \sim \text{Uniform} \binom{p}{k}$
- ▶ Measurement matrix  $\mathbf{X} \sim \prod_{i=1}^n \prod_{j=1}^p P_X(x_{i,j})$
- ▶ Non-zero entries  $\beta_S \sim P_{\beta_S}$
- ▶ Observations  $(\mathbf{Y}|\mathbf{X}, \beta) \sim P_{\mathbf{Y}|\mathbf{X}_S \beta_S}$

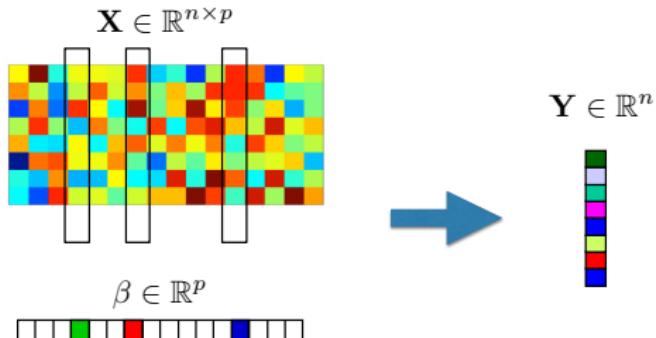
- ▶ Support recovery

$$P_e := \mathbb{P}[\hat{S} \neq S]$$

- ▶ Partial recovery

$$P_e(d_{\max}) := \mathbb{P}\left[|S \setminus \hat{S}| > d_{\max} \cup |\hat{S} \setminus S| > d_{\max}\right]$$

# More General Support Recovery



- ▶ General probabilistic models

- ▶ Support  $S \sim \text{Uniform} \binom{p}{k}$
- ▶ Measurement matrix  $\mathbf{X} \sim \prod_{i=1}^n \prod_{j=1}^p P_X(x_{i,j})$
- ▶ Non-zero entries  $\beta_S \sim P_{\beta_S}$
- ▶ Observations  $(\mathbf{Y}|\mathbf{X}, \beta) \sim P_{\mathbf{Y}|\mathbf{X}_S \beta_S}$

- ▶ Support recovery

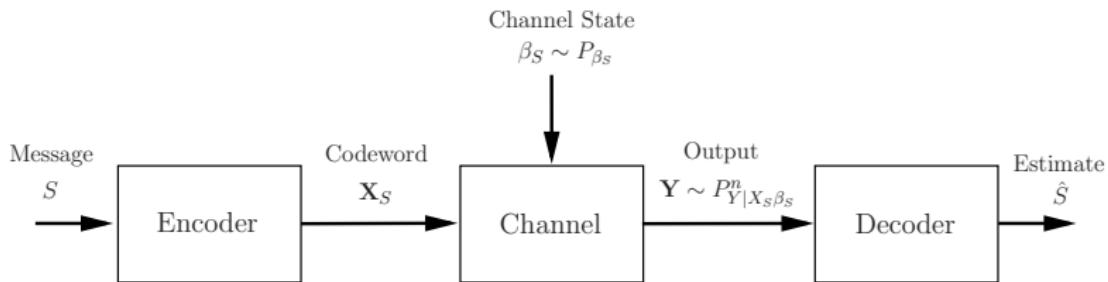
$$P_e := \mathbb{P}[\hat{S} \neq S]$$

- ▶ Partial recovery

$$P_e(d_{\max}) := \mathbb{P}\left[|S \setminus \hat{S}| > d_{\max} \cup |\hat{S} \setminus S| > d_{\max}\right]$$

- ▶ Goal: Conditions on  $n$  for  $P_e \rightarrow 0$  or  $P_e(d_{\max}) \rightarrow 0$

# Channel coding



e.g. see [Wainwright, 2009], [Atia and Saligrama, 2012], [Aksoylar et al., 2013]

## More General Support Recovery

- ▶ Steps for applying generalized non-asymptotic bounds:
  1. Construct a “typical set”  $\mathcal{T}_\beta$  of non-zero entries  $\beta_S$

# More General Support Recovery

- ▶ Steps for applying generalized non-asymptotic bounds:
  1. Construct a “typical set”  $\mathcal{T}_\beta$  of non-zero entries  $\beta_S$
  2. Bound the tail probabilities  $\mathbb{P}\left[\iota^n(\mathbf{X}_{s_{\text{dif}}}; \mathbf{Y} | \mathbf{X}_{s_{\text{seq}}}, \beta_s) \leq \mathbb{E}[\iota^n] \pm n\delta \mid \beta_s = b_s\right]$

## More General Support Recovery

- ▶ Steps for applying generalized non-asymptotic bounds:
  1. Construct a “typical set”  $\mathcal{T}_\beta$  of non-zero entries  $\beta_S$
  2. Bound the tail probabilities  $\mathbb{P}\left[\iota^n(\mathbf{X}_{s_{\text{dif}}}; \mathbf{Y} | \mathbf{X}_{s_{\text{seq}}}, \beta_s) \leq \mathbb{E}[\iota^n] \pm n\delta \mid \beta_s = b_s\right]$
  3. Control and simplify the remainder terms

# More General Support Recovery

- ▶ Steps for applying generalized non-asymptotic bounds:
  1. Construct a “typical set”  $\mathcal{T}_\beta$  of non-zero entries  $\beta_S$
  2. Bound the tail probabilities  $\mathbb{P}\left[\iota^n(\mathbf{X}_{s_{\text{dif}}}; \mathbf{Y} | \mathbf{X}_{s_{\text{seq}}}, \beta_s) \leq \mathbb{E}[\iota^n] \pm n\delta \mid \beta_s = b_s\right]$
  3. Control and simplify the remainder terms
- ▶ General form of corollaries:  $P_e \rightarrow 0$  if

$$n \geq \max_{(s_{\text{dif}}, s_{\text{seq}})} \frac{\log \binom{p-k}{|s_{\text{dif}}|}}{I(X_{s_{\text{dif}}}; Y | X_{s_{\text{seq}}}, \beta_s = b_s)} (1 + \eta) \quad \text{for all } b_s \in \mathcal{T}_\beta$$

and  $P_e \rightarrow 1$  if

$$n \leq \max_{(s_{\text{dif}}, s_{\text{seq}})} \frac{\log \binom{p-k+|s_{\text{dif}}|}{|s_{\text{dif}}|}}{I(X_{s_{\text{dif}}}; Y | X_{s_{\text{seq}}}, \beta_s = b_s)} (1 - \eta) \quad \text{for all } b_s \in \mathcal{T}_\beta$$

# Exact Recovery for Linear and 1-bit Models

Observation models

$$Y = \langle X, \beta \rangle + Z$$

$$Y = \text{sign}(\langle X, \beta \rangle + Z)$$

# Exact Recovery for Linear and 1-bit Models

Observation models

$$Y = \langle X, \beta \rangle + Z$$

$$Y = \text{sign}(\langle X, \beta \rangle + Z)$$

- ▶ Case 1:
  - ▶ Gaussian  $X$ , discrete  $\beta_S$ , sparsity  $k = \Theta(1)$ , low SNR
  - ▶ Necessary and sufficient conditions:

Linear	1-bit
$\max_{(s_{\text{dif}}, s_{\text{eq}})} \frac{ s_{\text{dif}}  \log p}{\frac{1}{2\sigma^2} \sum_{i \in s_{\text{dif}}} b_i^2}$	$\max_{(s_{\text{dif}}, s_{\text{eq}})} \frac{ s_{\text{dif}}  \log p}{\frac{1}{\pi\sigma^2} \sum_{i \in s_{\text{dif}}} b_i^2}$

- ▶ Only factor  $\frac{\pi}{2}$  difference

# Exact Recovery for Linear and 1-bit Models

Observation models

$$Y = \langle X, \beta \rangle + Z$$

$$Y = \text{sign}(\langle X, \beta \rangle + Z)$$

► Case 1:

- Gaussian  $X$ , discrete  $\beta_S$ , sparsity  $k = \Theta(1)$ , low SNR
- Necessary and sufficient conditions:

Linear	1-bit
$\max_{(s_{\text{dif}}, s_{\text{eq}})} \frac{ s_{\text{dif}}  \log p}{\frac{1}{2\sigma^2} \sum_{i \in s_{\text{dif}}} b_i^2}$	$\max_{(s_{\text{dif}}, s_{\text{eq}})} \frac{ s_{\text{dif}}  \log p}{\frac{1}{\pi\sigma^2} \sum_{i \in s_{\text{dif}}} b_i^2}$

► Only factor  $\frac{\pi}{2}$  difference

► Case 2:

- Gaussian  $X$ , fixed  $\beta_S$ , sparsity  $k = \Theta(p)$ , moderate SNR
- Conditions:

Linear	1-bit
$\Theta(p)$ sufficient	$\Omega(p \sqrt{\log p})$ necessary

# Partial Recovery for Linear and 1-bit Models

- ▶ Partial recovery

$$P_e(d_{\max}) := \mathbb{P}\left[|S \setminus \hat{S}| > d_{\max} \cup |\hat{S} \setminus S| > d_{\max}\right]$$

- ▶ Linear and 1-bit models
- ▶ Gaussian  $X$ , Gaussian  $\beta_S$ , sparsity  $k = o(p)$ , allowed errors  $d_{\max} = \lfloor \alpha^* k \rfloor$

# Partial Recovery for Linear and 1-bit Models

- ▶ Partial recovery

$$P_e(d_{\max}) := \mathbb{P}\left[|S \setminus \hat{S}| > d_{\max} \cup |\hat{S} \setminus S| > d_{\max}\right]$$

- ▶ Linear and 1-bit models
- ▶ Gaussian  $X$ , Gaussian  $\beta_S$ , sparsity  $k = o(p)$ , allowed errors  $d_{\max} = \lfloor \alpha^* k \rfloor$

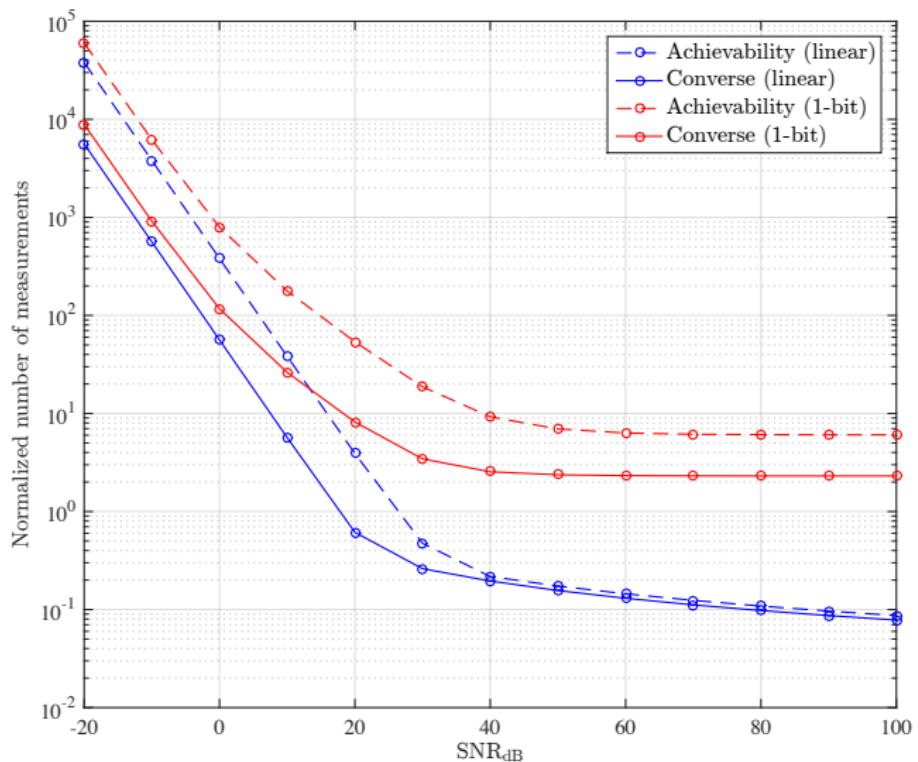
- ▶ Sufficient for  $P_e \rightarrow 0$ :

$$n \geq \max_{\alpha \in [\alpha^*, 1]} \frac{\alpha k \log \frac{p}{k}}{f(\alpha)} (1 + \eta)$$

- ▶ Necessary for  $P_e \not\rightarrow 1$ :

$$n \geq \max_{\alpha \in [\alpha^*, 1]} \frac{(\alpha - \alpha^*) k \log \frac{p}{k}}{f(\alpha)} (1 - \eta)$$

# Partial Recovery



# Conclusion

- ▶ Contributions:
  - ▶ New information-theoretic limits for group testing and other support recovery problems
  - ▶ Exact thresholds (phase transitions), or near-exact

# Conclusion

- ▶ Contributions:
  - ▶ New information-theoretic limits for group testing and other support recovery problems
  - ▶ Exact thresholds (phase transitions), or near-exact
- ▶ Implications for group testing:
  - ▶ Optimality of i.i.d. Bernoulli measurements in many cases
  - ▶ Limited gain by moving to partial recovery

# Conclusion

- ▶ Contributions:
  - ▶ New information-theoretic limits for group testing and other support recovery problems
  - ▶ Exact thresholds (phase transitions), or near-exact
- ▶ Implications for group testing:
  - ▶ Optimality of i.i.d. Bernoulli measurements in many cases
  - ▶ Limited gain by moving to partial recovery
- ▶ Future work:
  - ▶ Closing the remaining gaps (better concentration inequalities?)
  - ▶ Non-i.i.d. measurements
  - ▶ Applications to other non-linear models
  - ▶ “Information-spectrum” type analysis of other statistical problems
  - ▶ ...

# Conclusion

- ▶ Contributions:
  - ▶ New information-theoretic limits for group testing and other support recovery problems
  - ▶ Exact thresholds (phase transitions), or near-exact
- ▶ Implications for group testing:
  - ▶ Optimality of i.i.d. Bernoulli measurements in many cases
  - ▶ Limited gain by moving to partial recovery
- ▶ Future work:
  - ▶ Closing the remaining gaps (better concentration inequalities?)
  - ▶ Non-i.i.d. measurements
  - ▶ Applications to other non-linear models
  - ▶ “Information-spectrum” type analysis of other statistical problems
  - ▶ ...

## Further Details

- ▶ Further details (group testing):

<http://infoscience.epfl.ch/record/206886>

(accepted to 2016 SODA conference)

- ▶ Further details (general models):

<http://arxiv.org/abs/1501.07440>

(submitted to IEEE Transactions on Information Theory)

## References |

- M. Wainwright, "Information-theoretic limits on sparsity recovery in the high-dimensional and noisy setting," *IEEE Trans. Inf. Theory*, vol. 55, no. 12, pp. 5728–5741, Dec. 2009.
- G. Atia and V. Saligrama, "Boolean compressed sensing and noisy group testing," *IEEE Trans. Inf. Theory*, vol. 58, no. 3, pp. 1880–1901, March 2012.
- C. Aksoylar, G. Atia, and V. Saligrama, "Sparse signal processing with linear and non-linear observations: A unified Shannon theoretic approach," April 2013, <http://arxiv.org/abs/1304.0682>.
- T. S. Han, *Information-Spectrum Methods in Information Theory*. Springer, 2003.